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Modelling Wages and Prices for the Smaller OECD Countries

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bу

Kenichi Kawasaki, Peter Hoeller and Pierre Poret

October 1990



ECONOMICS AND STATISTICS DEPARTMENT

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This paper presents work on wage/price blocks for the smaller OECD countries which has been implemented in OECD's world econometric model, INTERLINK. The paper discusses theoretical, statistical and practical aspects of the estimation of business sector wage equations and five domestic demand deflators. It also presents a variety of diagnostic simulations in order to evaluate overall model properties.

* * * * *

Ce papier présente le bloc prix-salaires des petits pays du modèle mondial de l'OCDE, INTERLINK. Le papier examine les aspects théoriques, statistiques et pratiques des équations, estimées pour le taux de salaire du secteur des entreprises et cinq déflateurs de la demande intérieure. Il présente aussi un ensemble de simulations destinées à évaluer les propriétés variantielles globales du modèle.

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by

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October 1990

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MODELLING WAGES AND PRICES FOR THE SMALLER OECD COUNTRIES

INTRODUCTION

Developments in wage and price inflation are of central importance in macroeconomic analysis and forecasting. A complete set of wage and price equations for the major seven OECD countries was estimated and implemented in OECD's international model INTERLINK some time ago (see Stiehler, 1987). This paper now reports similar work for the smaller OECD countries.

The estimation of wage equations for the smaller countries adopts the same approach as for the large countries. The specification follows a standard approach to wage-setting where the long-run equilibrium for real wages is reached when labour markets operate at the "natural rate of unemployment". Disequilibrium forces include deviations of actual from trend productivity, terms-of-trade effects and sluggish adjustment of wages to recent price developments. As for prices, domestic demand deflators are modelled as a mark-up over unit costs with mark-ups being allowed to fluctuate cyclically. In the context of the INTERLINK model monetary policy influences inflation through various channels. There are relatively direct effects via exchange rates and import prices and less direct effects which work through changes in labour and goods market tightness.

Part I of this paper presents the stylised facts about the wage/price block and its relation to other parts of the model, especially monetary policy instruments. Part II discusses the specification for the wage and price equations and presents estimation results. It also includes wage/price block simulations. Part III presents full model simulation results. The final section includes some suggestions for future work.

I. THE STYLISED MODEL OF WAGE AND PRICE BEHAVIOUR

At the centre of the wage/price block is a standard theoretical model of wage determination: the rate of change of wages (w) is indexed to past (p) and expected (p_e) inflation and productivity growth in the short (q) and long run (q_n). The effects of labour-market disequilibria on wage growth are captured by the deviation of the rate of unemployment (U) from the natural rate (U_n). Furthermore, wage rates are assumed to react to differences in consumer (p) and producer price inflation (pgdp).

[1]
$$w = a_1p + (1-a_1)p_e + a_2(U-U_n) + a_3q + (1-a_3)q_n + a_4(p-pgdp)$$

Unitary elasticities of wages with respect to prices and productivity ensure first that the wage equation is homogeneous in price inflation and, second, that real wages will grow in line with labour productivity in the absence of a labour-market gap and changes in consumer relative to output prices. The wage equation could also be estimated in levels using an error-correction term. This would ensure, in addition, that income shares would return to "normal" levels in long-run equilibrium.

Assuming constant returns to scale and oligopolistic price-setting, changes in domestic demand deflators (p) are determined as a mark-up over unit cost changes in the long run. Cost elements include smoothed changes in unit labour costs (lc), capital costs (cc) and import costs (pm). Profit margins are allowed to depend positively on the rate of capacity utilisation (g), reflecting, for instance, the difference in the development of average and marginal cost over the cycle:

[2]
$$p = a_1 lc + a_2 cc + (1-a_1-a_2)pm + a_3g$$

Price inflation is stable if unit costs grow at a constant rate and excess demand is zero. An error-correction term between price and cost levels could be added to the specification; this would ensure constant profit margins in the long run. An important characteristic of the wage/price block is the

inter-connection of equilibria in the goods and labour markets. As capacity utilisation in the goods market is measured using maximum employment given by the "natural" rate as derived from the wage equation estimates, stable wage and price inflation is achieved only if both markets are in equilibrium at the same time (Torres and Martin, 1990).

Monetary policy is important in shaping inflationary developments, but monetary aggregates do not directly affect wage and price developments. In the context of the INTERLINK model, expansionary monetary policy leads to a reduction in nominal, and, given sluggish price adjustment, real short-term and long-term interest rates. Real interest rates enter the consumption and housing equations and alter demand for labour and capital inputs. In addition, interest-rate changes affect exchange rates, which in turn move foreign demand and trade prices. In sum, monetary policy affects demand deflators via exchange rates, import prices and capital costs and works indirectly through changes in the unemployment rate and excess demand in goods markets. An overview of the interaction between the different blocks of INTERLINK involved in this process is given in Richardson (1987 and 1988).

Whether simple models of aggregate wage and price behaviour as outlined above are an accurate description of major aspects of macro-relationships and of use for policy analysis is subject of recurrent controversy. If expectations are formed rationally, no policy trade-offs exist, even in the short run, as agents fully take account of monetary policy. Only imperfect information and adjustment costs can cause deviations from equilibrium. Furthermore, use of estimated equations for policy analysis will be faulty if expectations are not independent of policy rules (Lucas, 1976). Although the new classical critique of the Phillips curve has been influential, research on the microeconomics of sluggish adjustment as optimising behaviour, has given new credence to the augmented Phillips-curve approach adopted in this paper (1).

II. WAGE AND PRICE EQUATION ESTIMATES

1. Testing for long-run relationships

Whether the wage and price equations should be specified in terms of levels or growth rates is an important modelling issue. The answer depends on the existence of a stable long-run "equilibrium" level relationship. If such an equilibrium exists, the dynamics of an equation specified in growth-rate terms will be misspecified because the information that the dependent variable tends to revert to some trend level in the long run will be lost. On the other hand, if no stable "normal" level can be found, an equation with an error-correction mechanism will lead to spurious estimates.

The existence of a long-run equilibrium relationship can be examined by testing for "stationarity" of the series -- a series being defined as stationary when innovations do not permanently affect the long-run value. This issue is examined using unit-root tests. In principle, the hypothesis that the so-called wage (price) target is non-stationary is tested, or more specifically, whether the target is correlated with its one-period-lagged value with a coefficient not inferior to unity. In the case of a unitary coefficient, any shock to the variable will persist indefinitely. Table I gives a formal definition of stationarity and a detailed description of the tests.

The wage target was defined as the ratio of real consumption wages to current labour efficiency, assuming a unitary elasticity of wages with respect to prices and productivity. This ratio differs somewhat from usual definitions of the labour share as consumer prices are used instead of the output deflator, non-wage labour costs are excluded and labour productivity is adjusted for the effect of capital-labour substitution. Except for Ireland and Switzerland the tests unambiguously fail to reject the null hypothesis of non-stationarity, thus suggesting that there is no stable long-run equilibrium in terms of levels and justifying a specification of the wage equations in growth rate terms (Table 1, line (1)).

The price target selected for the test of mean reversion of profit margins is the ratio of the private consumption deflator to total costs (a weighted average of labour, capital and import costs). These cost components form the basis of the mark-up setting process. Tests suggest that the mark-up is stationary for Denmark, Finland, the Netherlands and Switzerland (Table 1, line (2)), thus supporting a specification in terms of levels for those countries. There is much less support in the data for stationarity of the mark-up term for the other countries. However, the cost weights have been imposed and a free estimation would have probably improved the stationarity of the mark-up term. It is also possible that the time to revert to trend exceeds the relatively short sample-period horizon used for these tests.

However, a feature of the results is that the price target is closer to stationarity than the wage target as the test statistics in Table 1 are generally higher for the former than for the latter. Thus, the inclusion of other costs than labour cost leads to a more stable mark-up than it would be if simply defined as a standard profit share. As longer-run changes in profit margins are also difficult to justify in simulations, the price equations were specified in levels.

2. Wage equation estimation results

The estimation of wage equations adopted here follows the approach summarised in Coe and Gagliardi (1985) and Coe (1985). That work focused on wage equations for the major seven OECD countries and four smaller OECD countries -- Australia, Austria, the Netherlands and Switzerland.

The dependent variable in the equations is the growth rate of business sector wages and salaries per employee (w) (2). It is related to a moving average of past changes in the private consumption deflator (p) and an activity variable (U). The activity variable is represented by the unemployment rate except for the case of Switzerland where a measure of employment is used. In addition, tests are carried out for short-run productivity (q) and producer price effects (p, pgdp). The natural rate of unemployment and trend productivity are not modelled explicitly (see below), but form part of the constant:

[3] $w = a_0 + a_1p - a_2U + a_3q + a_4(p-pgdp)$

The equations chosen for implementation in INTERLINK are presented in Table 2. The selection criteria used included goodness of fit, recent tracking performance and the overall simulation properties of the country model. A variety of specification and other issues are reported below.

Specification of the unemployment term. Linearity of the activity term was examined for each country. A linear specification implies that unemployment has the same impact on wage increases regardless of its level, while in the inverse case the effect of unemployment on wage increases is lower the higher the rate of unemployment. The use of the log of the unemployment rate implies that a given percentage change in the unemployment rate has the same effect regardless of its level. A linear version of the activity variable was chosen for Australia, Belgium, Greece, New Zealand and Switzerland, the inverse for Norway and Sweden and the logarithmic version for Austria, Denmark, Finland, Ireland, the Netherlands and Spain (3).

<u>Inflation</u>. Over a longer stretch of time wages are likely to follow inflation rates closely -- otherwise real wages would change indefinitely in response to changes in inflation. Coe (1985) showed that the estimated coefficients were insignificantly different from unity for the eleven countries in his sample. As homogeneity is also an important model property for large nominal shocks, a unitary coefficient of wage inflation with respect to the moving average of past changes in the consumption price deflator was imposed. The only exceptions are Australia and Spain, although homogeneity will be imposed for these two countries in the next round of estimation.

Adaptive expectations are assumed (4). Coe (1985) tested several forms of forward-looking expectations formation and arrived at the conclusion that such a specification did not improve estimates over specifications using current and past inflation (5). Furthermore, model-consistent expectations can not yet be generated in INTERLINK.

The role of producer prices. An output price effect, in the form of differences in growth rates between the business sector output and consumption

deflators has also been tested. While wage rates relative to consumer prices are relevant for labour supply, total compensation deflated by output prices is important for labour demand of employers. The weights between consumption and output price deflators are likely to depend on the slopes of the labour demand and supply curves, but could also reflect the fact that bargaining takes account of profitability in the business sector. The output price effect is large in Austria, Switzerland and New Zealand and still sizeable in Belgium, Greece, Ireland and Norway. Following an approach more akin to the Scandinavian model of wage determination, the output price deflator is replaced by the export price deflator in the case of Sweden (6).

<u>Productivity effects</u>. Current and lagged effects of business sector productivity growth were tested for all countries and, where significant, retained in the equations. Cyclical movements in productivity were found to have a significant effect in Belgium, Denmark, Finland, Greece, New Zealand, Norway, Spain and Switzerland. The short-run productivity effect captures the cyclical variability of bonus and overtime payments. In some countries the results are consistent with forms of bargaining behaviour which take explicit account of recent productivity developments.

The "natural" rate of unemployment. An important question for macroeconomic policy is whether current labour-market conditions are likely to lead to wage moderation or accelerating wage inflation. As argued above (equation [1]), this will largely depend on the difference between the current rate of unemployment and the natural rate of unemployment. In steady state, output and consumer prices and wages grow at the same rate, the unemployment rate is at its natural rate and productivity grows at its trend rate. Taking these conditions and rearranging equation [3] gives the following expression for the natural rate (U^*) :

- [4a] $U^* = (a_0 (1-a_3)q_n)/a_2$ for the linear case;
- [4b] $U^* = \exp[(a_0 (1-a_3)q_n)/a_2]$ for the logarithmic case;
- [4c] $U^* = a_2[(1-a_3)q_n a_0]$ for the inverse specification.

Natural rates implied by the Phillips-curve estimates are shown in Table 3. As the calculation of the natural rate depends on the development of

trend labour productivity growth, the numbers are shown for sub-periods. most of the smaller countries, natural rates have drifted upwards since the 1960s as labour productivity growth has slowed down. After a period in which actual unemployment rates have been much higher than natural rates in the early a period coinciding with wage moderation, unemployment rates are now to the estimated natural rates in Austria, Sweden, Finland and In some countries, for instance Spain and Ireland, current Switzerland. unemployment rates are still far above the estimated natural rate (7). needs to be stressed that such estimates of the natural rate provide only a rough guide as to when inflationary pressures stemming from the labour market may arise. Large confidence intervals around these estimates are likely to apply, due to imprecise coefficient estimates and potential misspecification. In particular, the calculations here assume that the only source of change of the natural rate stems from changes in trend productivity. Hence the growth of minimum wages, changes in the unemployment benefit replacement ratio and in unionisation, all factors which are thought to affect the natural rate of unemployment have not been taken into account.

3. Price equation estimation results

Equations for five domestic demand deflators have been estimated: the deflator for private consumption, for government non-wage consumption, for government investment, for business non-residential investment and for housing investment. Deflators for exports and imports had been estimated earlier (Herd, 1987). The deflator for inventory changes is exogenous for the smaller country models while final and total domestic demand, GDP and business sector output deflators are calculated by identity.

There is a rich choice of right-hand-side variables and specifications. The initial work on the price blocks was therefore confined to testing specifications for a limited number of countries. The choices reviewed and the findings of this experimental stage are summarised below without showing estimation results:

a) <u>Indirect taxes</u>. Proper modelling of price developments should rely on time series of net indirect tax rates by demand component. These are

available only for Denmark. For the other countries, information is available only for individual years from input-output tables, which are often out of date and not aggregated properly from the point of view of INTERLINK. Aggregate net indirect tax rates were therefore used for those deflators where the amount of taxation was significant. For instance, in countries with a VAT system the aggregate indirect tax rate for the private consumption deflator was used, but no indirect tax rate was applied to the business investment deflator. These assumptions, based on the available information about tax systems and input-output tables, are shown in Table 4.

- b) <u>Domestic costs</u>. Labour and capital cost are scaled by potential business sector output rather than actual output so as to smooth cyclical cost developments. The effect of real interest rates on capital costs has been reduced from its standard INTERLINK definition in order to better reflect the financial capital/output ratio. While labour and capital cost shares in the price equations were initially estimated, this often yielded unsatisfactory results. Large capital shares, for instance, produce counter-intuitive results in interest-rate simulations. The shares were therefore imposed at their historical values in the follow-up work.
- c) Import costs. Import costs have not been split into components, for instance, into energy and non-energy parts. Experimentation showed that this would affect estimation results little. Among large energy producers, a domestic energy component is implemented only for the Netherlands, since no suitable data are available for Norway. Experiments with moving weights between import and domestic costs were also made. As sample periods range back only to the early 1970s, using moving weights changed results little and they were thus dropped from the equations.
- d) Output gap term. For the smaller countries two gap terms are available from the INTERLINK supply block: one is the ratio of actual to "potential" output and the other the ratio of actual to "normal" output (Torres et al., 1989). The measure relating actual to "potential" output has been preferred because this measure reinforces the link between output and labour markets, as "potential" is evaluated at the natural rate of unemployment

implicit in the wage equations. It was thus used for all the other countries. As the coefficient on the gap term was incorrectly signed or of implausible size in estimated equations for some demand deflators, it was restricted to be the same across demand deflators.

e) Error-correction term. Error correction between domestic prices and cost components might be restricted to import costs only, establishing a law of one price in the long-run, in domestic cost only, or in a mixture of both. An error-correction term in both domestic and import costs was finally selected. As the freely-estimated share of import cost was often low, especially for the business investment deflator, this coefficient was imposed using input-output table information for the different deflators. Where such information was not available, a rough estimate was made, using aggregate import shares and deflator-specific information for similar countries.

The specification chosen and estimation results. With little variation across countries, the following equations were estimated as a simultaneous equation system. Small letters indicate growth rates of a variable and capitals indicate series in level or index form:

[5]
$$p = a_0 + (1-a_1) (a_2 lc + (1-a_2) cc) + a_1 pm + a_3 G$$

 $+ a_4 ((1-a_5) ln(a_2 lc(-1) + (1-a_2) cc(-1))/P(-1) + a_5 lnPM(-1)/P(-1)))$

where $P = P_{i+}/T$, the five demand deflators excluding indirect taxes

Pit = demand deflators including indirect taxes

T = tax base/(tax base - indirect taxes + subsidies)

LC = WSSE*(ETBPT/QBVPT), unit labour cost

WSSE = compensation per employee

ETBPT = potential business sector employment

QBVPT = potential business sector output

CC = PIB*(XRHOR + XRSCRB + 0.3*IRLRE)/100*(KBV/QBVPT), capital cost

PIB = business sector investment deflator

IRLRE = real interest rate

KBV = business sector capital stock

XRHOR, XRSCRB = parameters

PM = aggregate import price deflator

G = gap variable

Growth rates for domestic and import costs are moving averages with different lag lengths between domestic and foreign costs. Demand deflators are from national sources except for the government non-wage deflator (PCGNW), which has been calculated assuming no increase in the productivity of government employees:

[6] PCGNW = (CG - CGW)/(CGV - CGW/WRG)

where CG = government consumption in nominal terms

CGW = government wage bill

CGV = government consumption in real terms

WRG = government wage rate

The government deflator in the price block is then calculated by identity using the endogenous wage component (8).

Estimation results for the domestic demand deflators are summarised in Table 5 (9). The sample periods, lag lengths for import cost and domestic cost variables and the imposed labour income share are shown in parentheses. The gap term is restricted to be the same across deflators and marked with "@" in Table 5. Imposed parameters are marked with "I".

The gap term was not significant for Belgium, Denmark, Finland, Ireland and New Zealand. A small gap coefficient was imposed for Denmark and the Belgian equation re-estimated relaxing the restriction on the gap coefficient. For the other countries the gap term coefficient lies in a range of 0.03 for Austria to 0.17 for Australia. While the small coefficients for some countries seem reasonable in the light of their openness and Australia's large coefficient seems reasonable in terms of its small share of imports, it is difficult to generalise further as many other factors are potentially important in influencing pricing decisions. For instance, industrial policies severely limiting competition were in place in New Zealand over most of the sample period (OECD, 1987).

Estimation results for the government non-wage deflator are generally the poorest and in some cases a large number of dummy variables was used to arrive at reasonable equations. Coefficients were also imposed in the few equations where the error-correction coefficient was wrongly signed or where the short-term import price coefficient was larger than that for the long-term cost share.

Wage/price block dynamics. The dynamics of the price response to different cost shocks depends on the pattern of lags in the cost variables, the share of domestic and foreign costs in total cost and the error-correction coefficient. Table 6 takes Belgium as an example and shows the response of the private consumption deflator to different shocks holding a selection of other model variables constant. An increase in the level of all cost components by 1 per cent, holding other variables fixed, would increase the price deflator by 0.4 per cent after one year. Prices go up somewhat more than the initial shock in the third year and full homogeneity between cost and price levels is reached after the seventh year. The second panel in Table 6 shows an ex-ante shock to wages, with only prices, wages; capital cost and import prices endogenous. wage increase would lead to a sharp acceleration in prices in the first three years after which price increases decelerate and finally stabilise after six The third panel simulates an import price shock, with wages, capital costs and prices endogenous. Domestic price deflators and wages regain half of the difference to foreign prices after about three years. Adjustment slows afterwards and full adjustment is reached after seven years. A one-period shock to the gap term, shown in the fourth panel, leads to an increase in Both wages and prices stabilise after prices and afterwards in wages. seven years somewhat above the baseline level.

III. SIMULATION PROPERTIES

A series of full-model shocks were simulated in order to highlight the interaction between the wage/price block and the other parts of the model. The simulations were carried out with the Spring 1990 version of INTERLINK, which includes not only new supply blocks for the smaller countries, but also new consumption functions and major changes to the government accounts. The

results of the following shocks in single country mode are summarised in Tables 7 to 9:

- -- a fiscal shock to government non-wage expenditure with fixed nominal interest and exchange rates;
- -- a monetary shock which lowers nominal interest rates with fixed real government expenditure and floating exchange rates;
- -- an exchange-rate shock with fixed nominal interest rates and fixed real government expenditure.

Fiscal shock. The fiscal shock increases total domestic demand by about the same amount as government spending in the short run (Table 7). But short-run effects on aggregate output and labour markets, which are the most important determinants for the ensuing wage/price dynamics, differ across countries: demand leakages via imports nearly offset the boost to domestic demand in Belgium, while the output multiplier is large in Australia, Finland and Spain. Wage and price reactions, therefore, also differ significantly across countries. In the following years, wage and price increases stimulate demand further through a fall in real interest rates, but also dampen demand because rising export prices erode competitiveness.

Under the assumption of a fixed nominal exchange rate the erosion of competitiveness is the major factor pushing demand back to baseline in the longer run. Wages and prices also react with some lag and fall back towards baseline values. As expected, given their generally greater openness, the reaction of wages and prices for the average of the smaller countries to the fiscal shock is smaller than that of the large countries (Table 10). Also the peak deviation of prices and wages from baseline occurs considerably sooner.

Monetary shock. A reduction in the short-term interest rate lowers the borrowing cost of consumers and producers, boosting private consumption, housing and business investment (Table 8). At the same time it affects prices more directly by inducing a fall in the exchange rate and lowering capital cost. The net effect is a small increase in wages and prices in the short run. In most countries, inflation accelerates due to an inflation/devaluation spiral and a considerable fall in real interest rates. Thus, growth rates of wages

and prices are in general still higher than in the baseline by the fifth year. In comparison with the major seven countries, larger foreign demand leakages are a major reason for smaller output and price effects (Table 10).

Exchange-rate shock. A devaluation raises prices directly in the short run via increased import costs (Table 9). Wages respond with some lag. In addition, the short-run gain in competitiveness translates into an increase in net exports, which is large in some of the more open smaller countries. This, in turn, reduces unemployment and goods market slack and gives an additional boost to wage and price inflation. After about four years, wages and prices have risen by about 10 per cent above baseline (in response to a devaluation of 10 per cent) but then continue to increase further to reach a peak of about 12 per cent after six years before falling back (average for the smaller countries in Table 10).

This overshooting is due mainly to the fall in real interest rates, which boosts interest-sensitive demand components. But import prices also increase by more than implied by the devaluation as import price behaviour is affected by domestic price developments. In the longer run wage and price inflation turn around as the loss in competitiveness leads to slack in goods and labour markets. As shown in Table 10, the short-run response of wages and prices of the smaller countries is somewhat stronger than in the large countries, while peak deviations from the baseline are considerably larger.

IV. CONCLUDING REMARKS

As the most important use of INTERLINK is the simulation of short to medium-run policy scenarios, the major aim of the work on wage and domestic price formation for the smaller OECD Member countries has been guided by the need to improve simulation properties of the model.

Work on wage equations follows closely the earlier work for the large countries. Testing of these equations has proceeded much further than for domestic demand deflator equations, which have only recently been implemented. Further work on smaller country wage equations is likely to take up issues

which have recently been broached for the large countries: a more explicit modelling of the natural rate and of price expectations (Poret, 1990).

There are a number of areas where the new price blocks may be improved. First, sample periods are short for most countries and an effort has to be made to push samples backwards. Second, the way short-cuts have been taken, for instance with respect to the calculation of the government consumption deflator or the incorporation of indirect taxes, may be reconsidered. And third, a number of restrictions have been imposed, which could presumably be relaxed for a large number of countries.

Simulation results for the wage/price block and the full smaller country models, which now bring together earlier supply and demand side work, are encouraging. They are still preliminary, however, as further work is envisaged to improve the fiscal and monetary blocks and household accounts.

NOTES

- 1. Major components of the research on wage and price stickiness comprise theories based on staggered contracts in labour markets, implicit contract models or the "menu cost" approach to price adjustment. Bruno (1989), Grandmont (1989) and Layard and Bean (1989) recently reviewed the development of these strands of research.
- 2. Because of data limitations the dependent variable is total business sector compensation per employee in the cases of Greece, Ireland and New Zealand.
- 3. A hysteresis specification, where the labour market pressure variable is represented by a moving average of the unemployment rate, had the same explanatory power as the logarithmic version in the case of Spain.
- In all countries wages are indexed to present and past inflation in a more or less formal way. However, indexation is usually not 100 per cent, nor does it cover all the labour force. In addition, multi-year contracts do not allow for a quick feed-through of inflation into wage increases.
- 5. However, recent studies (Englander and Stone, 1989; Moghadam and Wren-Lewis, 1989; and Poret, 1990) found opposite results.
- 6. An extended Scandinavian model of inflation is estimated for Sweden in a recent OECD Economic Survey for Sweden (OECD, 1989).
- 7. The average natural rate (\overline{U}^*) is close to the average unemployment rate (U) over the sample period for the linear case. As the equation passes through the mean values of the variables, rearrangement of the wage equation gives:

$$\overline{w} = a_0 + \overline{p} + a_3 \overline{q} - a_2 \overline{U}$$
 and

$$\overline{U} = (a_0 - (\overline{w} - \overline{p}) + a_3\overline{q})/a_2.$$

If real wage growth equals productivity in the long run, the second equation is the same as the one used for calculating the natural rate in the linear case and $\vec{U} = \vec{U}^*$. In the logarithmic case, U^* is equal to the geometric average of the actual unemployment rate and in the inverse case to the harmonic average. As both the geometric and harmonic averages are smaller than the arithmetic average, natural rates in the non-linear cases may appear low as compared to average actual unemployment rates.

8. Growth in government wage rates depend on private sector wage inflation in nearly all small country models. Estimated equations only exist for Denmark, the Netherlands and Sweden, where government wages are related to a distributed lag of private sector wage rates and an error-correction term in the levels of private and government sector wages.

9. The private consumption deflator for Australia also includes a term for administered prices, with a weight of 0.27. In simulations the administered component is linked to the overall private consumption deflator. In the case of the Netherlands, the private consumption deflator includes an additional component for quaternary services with a weight of 0.28 and for natural gas deliveries with a weight of 0.07. The latter component also enters the government non-wage deflator with a weight of 0.06.

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Table 1
Unit root tests

			Alternat	ive 3	Alternative 2	Alternative 1			
			Φ ₃	Φ2	Φ ₁	^t α'			
Australia	(35)	(1)	2.24	1.58	1.24	0.39			
•		(2)	2.48	2.25	2.24	-1.86			
Austria	(41)	(1)	1.61	1.10	1.10	0.23			
		(2)	2.24	2.30	1.98	-1.72			
Belgium	(34)	(1)	6.55	4.36	1.43	-0.15			
_		(2)				-1.59			
Denmark	(34)	(1)	3.64	2.64	0.47	0.70			
		(2)	7.64*		-				
Finland	(39)	(1)	4.75	3.18	3.30	-0.26			
		(2)	11.68*						
Greece	(40)	(1)	3.52	2.35	1.69	-0.51			
	•	(2)	3.87	3.42	1.47	0.39			
Ireland	(35)	(1)	7.11*		***				
		(2)	5.04	4.04	1.10	-0.73			
Netherlands	(33)	(1)	2.10	1.93	1.08	1.12			
		(2)	2.48	1.88	3.58	-2.19*			
New Zealand	(41)	(1)	2.56	1.71	1.92	-0.36			
4		(2)	1.72	1.15	1.74	-1.05			
Norway	(40)	(1)	2.87	2.17	0.56	0.70			
		(2)	2.24	2.15	2.17	-1.79			
Spain	(35)	(1)			e	1.32			
_		(2)	3.41	2.44	2.77	-1.52			
Sweden	(35)	(1)	1.86	1.69	1.59	0.94			
	-	(2)	2.36	1.78	1.12	-1.12			
Switzerland	(40)	(1)	3.42	3.86	6.95*				
		(2)	10.80*						

In brackets: number (T) of observations (semi-annual data). An asterisk denotes that the non-stationarity hypothesis is rejected; that is, the series are stationary.

⁽¹⁾ Logarithm of the ratio of real consumption wages to current labour efficiency (business sector).

⁽²⁾ Logarithm of the ratio of the private consumption deflator to total weighted costs.

Note to Table 1

Description of the test

A variable, χ , is said to be stationary if the coefficient (ρ) of the one-period lagged χ is inferior to unity in an equation such as: $\chi_{\mathsf{t}} = \mu + \rho \chi_{\mathsf{t}-1}$, or equivalently if $\alpha = \rho$ -1 is strictly inferior to 0 in $\Delta \chi_{\mathsf{t}} = \mu + \alpha \chi_{\mathsf{t}-1}$.

More specifically, the test is the Augmented Dickey-Fuller test (using fourth-order correction). The testing strategy (Perron, <u>Journal of Economic Dynamics and Control</u> (12), 1988) goes from "general" to "specific". One starts by testing the null hypothesis of a unit root with a drift and no time trend $(\Delta \chi_t = \mu + \alpha \chi_{t-1} + \epsilon_t)$ against Alternative 3, which is:

$$\Delta \chi_{t} = \mu + \beta(t-T/2) + \alpha \chi_{t-1} + \sum_{i=1}^{4} \gamma_{i} \Delta \chi_{t-i} + u_{t}.$$

The test statistic (Φ_3) is therefore for the constraint that $\alpha=\beta=0$. If Φ_3 exceeds the critical value, one rejects the null (i.e. "acccepts" that χ is stationary, perhaps with a time trend; the hypothesis that $\alpha=0$ and β is non-zero is not tested) and the procedure ends. If not, one would "accept" that χ is a random walk, but this result could be due to the fact that the assumption of a non-zero drift in the null is erroneous. To check that, one repeats the process, except the null has no drift $(\Delta\chi_t=\epsilon_t)$. The test statistic (Φ_2) is thus for the constraint that $\mu=\alpha=\beta=0$. If Φ_2 exceeds the critical value, this is tantamount to rejecting $\mu=0$ (given the previous test). Thus, the original null was fine, and the procedure ends by "accepting" non-stationarity on the basis of Φ_3 . If Φ_2 does not exceed the critical value, the appropriate null does not include a drift. However, the procedure cannot end by acceptance of driftless non-stationarity because Φ_2 is not invariant with respect to the presence of a time trend in the alternative. Therefore, a more appropriate test may be against Alternative 2, which is:

$$\Delta \chi_{\mathsf{t}} = \mu^* + \alpha^* \chi_{\mathsf{t-1}} + \sum_{\mathsf{i=1}}^{\mathsf{r}} \gamma_{\mathsf{i}}^* \Delta \chi_{\mathsf{t-i}} + u_{\mathsf{t}}^*.$$

The test statistic (Φ_1) is thus for the constraint that $\mu^*=\alpha^*=0$ (but under a different null than Φ_3). If Φ_1 exceeds the critical value, one rejects the null hypothesis of non-stationarity and the procedure ends, Φ_1 being invariant to the presence of a non-zero mean in the alternative. If Φ_1 is below the critical value, one accepts the non stationarity after having checked, on the basis of its t-statistics, that α is not significantly less than zero in alternative 1 (where $\mu=0$, as the series may have a zero mean in the alternative):

$$\Delta \chi_{t} = \alpha' \chi_{t-1} + \sum_{i=1}^{4} \gamma_{i}' \Delta \chi_{t-i} + u_{t}'.$$

The column t_{α} , gives the t-statistics associated with the coefficient of χ_{t-1} (α') in Alternative 1. Φ_3 , Φ_2 , Φ_1 are the F-statistics associated with the joint test of the null hypotheses, respectively (μ , β , α) = (μ , 0, 0); (μ *, α *) = (0, 0, 0); (μ *, α *) = (0, 0). Critical values do not follow the usual t-and F-distributions. For 50 observations, the critical value for the t-statistics at the 5 per cent level for Alternative 1 is -1.95 (Fuller (1976), Introduction to Statistical Time Series, μ . 373). The critical values for Φ_3 , Φ_2 and Φ_1 are 6.73, 5.13, 4.86 (Dickey and Fuller (1981), Econometrica 4, μ . 1063).

The above testing strategy ensures that the null hypothesis of non-stationarity with a non-zero drift is checked first. Φ tests are not reported when α is positive (that is, $\rho>1$) since, although α may be significantly different from zero, the series is obviously non-stationary (it follows an explosive process).

Table 2
Aggregate wage equations (a)

	Constant	Onempi	oyment re	(0)	Inflation	Productivity	Output price	SEE	DW	A djR ²
		ט	lnU	1/0	(b)	growth (c)	effect (d)			
Australia	2.11	-0.39			1.14			1.94	1.9	0.66
197011-198511	(0.8)	(2.6)			(2.8)					
Austria	2.10		-1.73		1		-1.03	1.45	2.6	0.50
19691-8611	(5.4)		(3.1)	ı	(Ŧ)		(1.8)			
Belgium	5,13	-0.54			1	0.76	-0.46	1.41	2.5	0.84
1971-88	(4.3)	(5.6)			(±)	(4.0)	(2.4)			
Denmark	1.96		-1.11		1	0.46		0.66	2.0	0.80
196611-8711	(3.5)		(5.6)		(±)	(2.1)				
Finland	3.20		-2.06		1	0.83		1.61	2.2	0.48
196711-8611	(2.1)		(2.9)		(±)	(2.4)				
Greece	2.10	-0.41			i	1.01	-0.50	2.17	1.6	0.47
19631-8611	(2.3)	(2.5)			(±)	(6.1)	(±)			
Íreland	6.40		-2.16		1		-0.73	2.17	1.4	0.49
19701-8611	(2.7)		(2.0)		(Ŧ)		(3.3)			
Netherlands	4.13		-1.74		1			0.85	2.2	0.75
19711-8511	(7.4)		(6.0)		(i)					
New Zealand	1.10	-0.65			1	0.78	-0.80	2.72	1.9	0.55
19671-8611	(1.7)	(2.4)		7	(±)	(1.8)	(1.8)			
Norway	-1.70			2.73	1	0.81	-0.60	1.61	2.2	0.45
19711-8611	(1.4)			(1.3)	(±)	(3.0)	(1.7)			
Spain	2.94		-1.74		1.03	0.77		1.78	2.3	0.58
19641-8411	(1.8)		(0.5)		(6.4)	(1.8)				
Sweden	-3.81			7.78	1	-	-0.62	1.62	1.3	0.76
1972II-84II	(2.6)			(2.5)	(T)		(3.3)	-		
Switzerland	38.10	0.39(•)		1	0.56	-0.91	0.86	1.6	0.76
196211-8611	(3.3)	(3.3)			(i)	(4.4)	(5.4)			

a) The dependent variable is the growth of the wage rate except for Greece, Ireland and New Zealand. All equations are estimated by two-stage least squares on seasonally-adjusted semi-annual data except for Belgium where annual data are used. T-statistics appear in parentheses below the coefficient estimates.

b) Inflation is defined as a moving average of current and past growth of the personal consumption deflator for all countries except Belgium for which no lagged effect of inflation was found. A two-period weighted moving average is used for Norway (weights of 0.39, 0.61) and Ireland (0.62, 0.38). For the other countries, the inflation term is a moving average of either two periods (Australia, the Netherlands, Switzerland, Greece), three periods (Austria, Finland) or four periods (Spain).

c) Productivity growth is specified as a two-period moving average for Norway; for Denmark, Finland and Spain, it is a three-period moving average; for Switzerland it is unlagged. Business sector productivity is used.

d) Difference between the growth of the private consumption deflator and the growth of the business sector output deflator. The export deflator for goods and services is used for Sweden, instead of the business sector output deflator.

e) The activity variable for Switzerland is a measure of the employment rate defined as total employment divided by a lagged two-period moving average of the labour force, multiplied by 100.

Note to Table 2

Country-specific dummy variables in the wage equations

Country	Values	Parameter estimate	t-ratio
Sweden	-1 and 1 in the first and second semesters of 1973, 1974 and 1979	3.34	(4.99)
	-1 and 1 in the first and second semesters of 1977	-5.92	(5.14)
	-1 and 1 in the first and second semesters of 1973, 1980, 1981	-3.14	(4.70)
New Zealand	1 in 1986 I	12.27	(3.1)
Australia	1 from 1974 I to 1974 II	5.81	(2.24)
Netherlands	1 from 1980 I to 1981 I	-3.48	(6.70)
Austria	1 in 1971 I	6.5	(4.2)

Table 3

Natural rate of unemployment estimates

	Unem	ployment ra	ite	Natural rate of unemployment					
	1966/87	7 1980/87 1988		1966/73	1974/79	1980/87			
Australia	4.9	7.7	7.1	2.8	3.2	5.1			
Austria	2.0	3.1	3.6	0.7	1.2	2.8			
Belgium	6.5	11.4	10.3	5.9	6.5	6.5			
Denmark	5.3	8.9	8.6	4.3	6.0	6.9			
Finland	3 7	5 1	4.6	4.0	4.0	4.1			
Greece	4.3	6.4	7.7	5.2	5.1	5.1			
Ireland	9.1	13.8	16 7	7.3	6.5	5.9			
Netherlands	4.3	8.7	8.3	2.5	4.3	6.0			
New Zealand	1.8	3.9	6.0	1.6	1.5	1.6			
Norway	1.9	2.4	3.2	1.2	1.5	1.5			
Spain	8.4	17.9	19.5	6.7	6.7	7.4			
Sweden	1.9	2.3	1.6	2.6	2.6	2.8			
Switzerland	0.3	0.6	0.7	0.5	1.3	0.8			

Table 4

Indirect tax rate assumptions
(Value is one where indirect tax rate applies; zero otherwise)

	PCP	PCGNW	PIH	PIB	PIG
Australia	1	0 `	0	0	0
Austria	1	1	1	0	1
Belgium	1	1	1	0	1
Denmark	1	1	1	0	1
Finland	1	0	0	0	0
Greece	1	0	0	0	0
Ireland	1	0	0	0	0
Netherlands	1	0	1	1	1
New Zealand	1	1	1	1	1
Norway	1	0	0	0	0
Spain	1	0	0	0	0
Sweden	1	1	1	0	0
Switzerland	1	1 .	1	1	1
			,		

PCP Private consumption deflator

PCGNW Government non-wage consumption deflator

PIH Residential construction deflator

PIB Business investment deflator

PIG Government investment deflator

Table 5
Equations for domestic demand deflators (a)

	Constant	Short-term import price effect	Error- correction term	Long-term import price effect	Gap term (b)	rmse	R ²	DW	Dh	RHO1
<u>Australia</u> (Sample - 19	731-198611;	lag = 1 for	import and 4	for domestic	cost; lab	our co	st she	re = 0	. 62)	
Private consumption deflator (c)	0.52 (1.2)	0.11 (2.3)	0.10 (1.4)	0.15 (I)	@0.17 (3.7)	1.07	0.67	1.82	:	
Government non-wage consumption deflator	1.72 (2.5)	0.07 (0.5)	0.62 (10.2)	0.10 (I)	@0.17 (3.7)	3.19	0.75	1.49	:	•
Residential con- struction deflator	0.96 (1.7)	0.11 (1.1)	0.16 (1.0)	0.15 (I)	@ 0.17 (3.7)	1.86	0.51	1.50	2.56	0.34 (2.0)
Business investment deflator	2.04 (5.5)	0.29 (5.3)	0.40 (5.9)	0. 4 5 (I)	@0.17 (3.7)	1.26	0.80	1.39		
Government investment deflator	0.30 (0.8)	0.28 (3.9)	0.13 (2.1)	0.30 (I)	@0.17 (3.7)	1.65	0.72	1.76		
Austria (Sample = 1971)	II-1986II;	lag = 1 for i	mport and 6	for domestic o	ost; labo	ur cos	t shex	·• = 0.	56)	
Private consumption deflator	0.85 (0.4)	0.14 (4.2)	0.57 (7.8)	0.25 (I)	@0.03 (1.3)	0.82	0.64	2.39	-1.15	0.92
Government non-wage consumption deflator	0.75 (1.1)	0.09 (0.6)	0.24 (4.4)	0,20 (I)	@0.03 (1.3)	2.59	0.86	177	0.55	0.30
Residential con- struction deflator	-0.43 (0.8)	0.16 (4.2)	0.21 (4.8)	0.20 (I)	@0.03 (1.3)	1.66	0.61	1.91	0.12	0.36 (6.2)
Business investment deflator	-0.06 (0.3)	0.25 (I)	0.19 (2.3)	0.45 (I)	@0.03 (1.3)	1.10	0.54	2.45		
Government investment deflator	-0.16 (0.4)	0.15 (I)	0.12 (4.4)	0.25 (I)	0 0.03 (1.3)	1.52	0.43	1.72		
Belgium (Sample - 1974)	I-1986II;	lag = 2 for in	mport and 6 f	for domestic co	ost; labou	ır cost	share	- 0.6	i9)	
Private consumption deflator	0.82 (3.2)	0.27 (8.1)	0.11 (5.4)	0.36 (I)	0.10 (2.5)	0.78	0.82	1.30		
Government non-wage consumption deflator	1.11 (2.3)	0.15 (1.4)	0.09 (1.6)	0.36 (I)	0.0 (I)	2.36	0.46	1.47		
Residential con- struction deflator	-0.22 (0.3)	0.28 (I)	0.0 (I)	0.28 (I)	0.0 (I)	1.24	0.78	1.65	0.67	0.76 (7.4
Business investment deflator	0.25 (0.3)	0.22 (3.5)	0.06 (1.7)	0.68 (I)	0.09 (0.8)	1.00	0.70	1.76	-0.04	.0.4 (3.5
Government investment deflator	-0.48 (2.1)	0.28 (I)	0.0 (I)	0.28 (I)	0.0 (I)	0.82	0.83	1.14	2.43	0.2 (2.9
Denmark (Sample - 1974)	I-1986II;	lag = 1 for in	mport and 3 f	or domestic co	ost; labou	ır cost	: share	- 0.6	56)	
Private consumption	0.24	0.18 (3.9)	0.34 (3.5)	0.24 (I)	@0.06 (I)	1.14	0.71	1.93		
Government non-wage consumption deflator	1.26 (2.2)	0.22 (2.0)	0.56 (3.3)	0.27 (I)	80.06 (I)	2.24	0.65	1.35		
Residential con- struction deflator	1.06 (1.0)	0.24 (I)	0.45 (1.1)	0.19 (I)	@ 0.06 (I)	0.83	0.91	1.43		0.7 (2.5
Business investment deflator	-0.03 (0.1)	0.30 (I)	0.06 (1.0)	0.59 (I)	8 0.06	0.85	0.80	1.65	2.99	0.5 (2.7
Government investment	0.19	0.25 (5.4)	0.25 (1.1)	0.20 (I)	@ 0.06 (I)	0.92	0.85	1.60		0.4

Table 5 (continued)

	Constant	Short-term import price effect	Error- correction term	Long-term import price effect	Gap term (b)	rmse	R ²	DW	Dh	RH01
inland (Sample = 1972I-	198611;	lag = 1 for im	port and 5 f	or domestic co	st; labou	r cost	share	- 0.7	0)	
rivate consumption eflator	-0.51 (2.0)	0.18 (4.7)	0.25 (4.5)	0.22 (I)		1.23	0.79	2.01		
Sovernment non-wage consumption deflator	-2.53 (3.3)	0.16 (2.0)	0.41 (4.7)	0.20 (I)		2.51	0.70	1.42		
tesidential con-	-5.12 (3.4)	0.11 (0.9)	0.83 (5.4)	0.20 (1)		3.61	0.52	1.12	2.83	0.42 (3.5)
dusiness investment deflator	0.12 (0.3)	0.25 (3.6)	0.06 (0.6)	0.50 (I)		2.00	0.63	2.70		
Government investment deflator	-0.77 (1.7)	0.28 (4.9)	0.27 (3.6)	0.20 (I)		1.66	0.74	1.34	2.76	0.10 (0.7)
Freece (Sample = 1971I-1	198611;	lag = 1 for imp	ort and 6 fo	r domestic cos	ıt; labour	cost	share	- 0.60))	
Private consumption leflator	-1.36 (1.1)	0.20 (I)	0.09 (1.1)	0.20 (1)	@0.10 (I)	2.90	0.50	1.40	4.21	0.34 (2.1)
Sovernment non-wage consumption deflator	-4.29 (1.1)	0.20 (I)	0.11 (0.9)	0.20 (I)	@0.10 (I)	2.76	0.67	1.56	15.19	0.58 (3.3)
esidential con- truction deflator	5.44 (1.4)	0.10 (I)	0.68 (6.4)	0.10 (1)	@ 0.10 (I)	2.88	0.55	1.38	1.72	0.87 (11.4)
usiness investment eflator	-0.38 (0.6)	0.35 (I)	0.15 (1.8)	0.67 (I)	@0.10 (I)	3.16	0.57	1.55		,
Sovernment investment deflator	1.29	0.10 (I)	0.0 (I)	0.10 (I)	@0.10 (I)	2.91	0.41	1.27		
<u> </u>	-1986II;	lag - 2 for is	mport and 6 i	for domestic co	ost; labou	ir cost	share	- 0.7	74)	
Private consumption	-0.11 (0.2)	0.25 (3.0)	0.08	0.35 (I)		1.87	0.67	1.62		
Sovernment non-wage	-0.46 (0.6)	0.13 (1.2)	0.11 (1.7)	0.30 (I)		2.27	0.64	1.26		
Residential con- struction deflator	-0.12 (0.1)	Ö.24 (1.6)	0.16 (1.4)	0.30 (I)		2.76	0.55	1.25	5.10	0.48 (2.6)
Business investment deflator	-0.23 (0.5)	0.47	0.01	0.60 (I)		2.20	0.71	2.62		
overnment investment	0.18	0.33 (4.0)	0.08	0.44 (I)		1.50	0.82	1.51	1.43	0.65 (4.7)
etherlands (Sample = 1	9741-1986	II; lag = 1 fc	or import and	d 6 for domest:	ic cost;	labour	cost s	hare .	- 0.64)	
Private consumption deflator (d)	-0.12 (0.3)	0.23 (I)	0.25 (5.9)	0.35 (I)	@0.09 (7.6)	1.40	0.24	1.50	0.61	0.27
Sovernment non-wage consumption deflator (d	-0.80	0.35	0.62 (5.7)	0.27 (I)	@0.09 (7.6)	1.26	0.77	1.32	1.05	0.36
COLEURDITON CETTATOR (C)	0.05	0.31	0.02	0.22	@0.09 (7.6)	1.89	0.30	1.25		, ,
Residential con-		/10 11	/* **							
Residential con- struction deflator Business investment deflator	(0.1) -0.14 (0.4)	(19.1) 0.31 (I)	(1.2) 0.24 (3.7)	(I) 0.47 (I)	(7.0) (0.09 (7.6)	1.70	0.30	1.38		

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Table 5 (continued)

	Constant	Short-term import price effect	Error- correction term	Long-term import price effect	Gap term (b)	RMSE	R ²	DW	Dh	ŘHO1
New Zealand (Sample -	19721-198611	; lag = 2 fo	r import and	6 for domesti	c cost;]	abour	cost s	hare -	0.52)	
Private consumption deflator	0.73 (1.4)	0.22 (3.1)	0.14 (2.1)	0.32 (I)		2.10	0.37	1.44		
Government non-wage consumption deflator	-1.26 (1.6)	0.34 (2.7)	0.10 (1.8)	0.32 (I)		3.73	0.28	1.57		
Residential con- struction deflator	1.01	0.21	0.0 (I)	0.24 (I)		2.57	0.35	1.49	1.90	0.36 (2.5)
Business investment deflator	-1.56 (1.6)	0.42	0.53 (2.7)	0.55 (I)		2.89	0.56	1.28		0.36
Government investment	-1.80	0.41	0.57	0.55		2.67	0.60	1.34		0.41
	(1.8)	(3.3)	(2.9)	(I)						(2.1)
Norway (Sample = 1971I				r domestic cos 0.30					')	
Private consumption deflator	0.39 (1.2)	0.25 (3.2)	0.25	(I)	@0.08 (1.8)	1.62	0.31	1.78		
Government non-wage consumption deflator	0.58 (0.3)	0.30 (I)	0.13 (1.2)	0.43 (I)	@0.08 (1.8)	3.36	0.81	1.58		0.43 (2.0)
Residential con- struction deflator	0.30 (0.7)	0.20 (I)	0.09	0.22 (I)	@0.08 (1.8)	1.55	0.42	1.27		
Business investment deflator	0.20	0.30 (I)	0.30 (3.4)	0.55 (I)	@0.08 (1.8)	2.56	0.70	1.94		
Government investment deflator	-0.27 (0.4)	0.20 (I)	0.03 (0.8)	0.22 (I)	6 0.08 (1.8)	1.46	0.38	1.31	5.21	0.11
<u>Spain</u> (Sample = 1974I-	1986II; lag	g = 1 for impo	rt and 5 for	domestic cost	; labour	cost i	share -	0.66)		
Private consumption deflator	-0.87 (1.4)	0.12 (2.3)	0.27 (3.2)	0.33 (I)	@0.06 (2.3)	2.47	0.50	2.36		
Government non-wage consumption deflator	-1.84 (2.5)	0.17 (1.9)	0.47 (4.9)	0.36 (I)	@ 0.06 (2.3)	3.59	0.42	1.26		
Residential con- struction deflator	-1.33 (2.3)	0.07 (2.6)	0.20 (6.0)	0.12 (I)	@0.06 (2.3)	1.52	0.84	0.99	2.74	0.25 (3.5)
Business investment deflator	-0.31 (0.8)	0.12 (2.9)	0.15 (7.8)	0.47 (I)	6 0.06 (2.3)	1.91	0.47	2.13		
Government investment deflator	-0.40 (1.3)	0.10 (3.1)	0.16 (6.0)	0.28 (I)	@ 0.06 (2.3)	1.43	0.68	1.92		
Sweden (Sample = 1973I	-19851I; la	ng = 1 for imp	ort and 6 fo	r domestic cos	ıt; labou	r cost	share	- 0.75	5)	
Private consumption deflator	0.79 (2.2)	0.24 (3.8)	0.09 (1.8)	0.25 (I)	@0.09 (2.2)	1.56	0.35	1.88		,
Government non-wage consumption deflator	1.53 (2.9)	0.25 (I)	0.21 (2.4)	0.29 (I)	@0.09 (2.2)	2.69	0.60	2.00		
Residential con- struction deflator	0.71 (0.4)	0.14 (3.2)	0.05 (0.8)	0.31 (I)	@ 0.09 (2.2)	1.05	0.82	1.57	1.20	0.87 (7.2)
Business investment deflator	0.12	0.32 (5.5)	0.03 (1.2)	0.31 (I)	@0.09 (2.2)	1.53	0.60	2.59		•
Government investment	0.48	0.18	0.04	0.31	@ 0.09	1.01	0.49	1.35		

Table 5 (continued)

	Constant	Short-term import price effect	Error- correction term	Long-term import price effect	Gap term	RMSE	R ²	DW	Dh	RH01
Switzerland (Sample =	19711-198611	; lag = 1 fo	r import and	4 for domesti	.c cost;	labour	cost	share -	0.67)	
Private consumption	0.15	0.17	0.0	0.25	60.11	0.94	0.70	2.01		
deflator	(0.9)	(4.6)	(I)	(I)	(3.9)					
Government non-wage	-1.96	0.20	0.12	0.25	0 0.11	2.23	0.57	2.19		
consumption deflator	(2.8)	(2.3)	(2.2)	(I)	(3.9)					
Residential con-	0.93	0.08	0.36	0.20	60.11	0.86	0.86	1.14	2.61	0.97
struction deflator	(1.7)	(4.0)	(4.2)	(I) ·	(3.9)					(11.6) -0.28 (3.5)
Business investment	-1.87	0.21	0.25	0.40	60.11	1.43	0.62	2.29		(3.3)
deflator	(5.0)	(3.8)	(4.6)	(I)	(3.9)					
Government investment	1.38	0.16	0.41	0.30	€0.11	1.59	0.62	1.05	3.34	0.14
deflator	(1.7)	(5.1)	(3.8)	(I)	(3.9)					(1.1) 0.44 (4.5)

a) The dependent variables are the growth in domestic demand deflators. The derivation of the government non-wage deflator is explained in the text. Equations are estimated by MINDIS on seasonally-adjusted semi-annual data. In a few countries, for instance in Belgium, semi-annual values are interpolated annual numbers. T-statistics appear in parentheses below coefficients. A coefficient with an (I) below is imposed, a coefficient with @ at the side is constrained across equations.

b) The gap variable is a two-semester moving average for all the countries.

c) Includes a term for an administered component of consumer prices with a weight of 0.27.

d) Includes a term for quaternary services and domestic gas deliveries with weights of 0.28 and 0.07 respectively in the private consumption deflator, and for domestic gas deliveries with a weight of 0.06 in the government non-wage consumption deflator.

e) RH02.

Note to Table 5 Country-specific dummy variables in the price equations (Values -- 1 unless otherwise indicated)

			Parameter	t-value
Australia	PCGNW	1973I; -1 in 1973II	-3.19	2.0
		1977I,1978I; -1 in 1977II,1978II	-2.87	2.7
		1981I; -1 in 1981II	-5.29	3.6
		1982I; -1 in 1982II	6.67	4.6
		1985I,1985II,1986I,1986II	-9.22	7.1
Austria	PCGNW			16.0
		1974I,1975I,1976I; -1 in 1974II,1975II,1976II		15.7
		1978I; -1 in 1978II	6.57	8.3
		1981I,1982I; -1 in 1981II,1982II	2.31	4.4
		1983I; -1 in 1983II	-1.98	2.5
		1985I; -1 in 1985II	0.82	1.1
	PIB	1971I,1972I; -1 in 1971II,1972II	2.99	7.2
Belgium	PCGNW	1974II,1975I; -1 in 1974I,1975II	2.52	2.7
Finland	PCGNW	1972I,1973I,1974I; -1 in 1972II,1973II,1974II 1975I,1976I,1977I,1978I; -1 in 1975II,1976II,	4.10	5.3
		1977II,1978II	-3.46	4.9
	PIH	1972I; -1 in 1972II	-2.95	2.7
		1975I; -1 in 1975II	-7.18	4.2
		1977I,1978I; -1 in 1977II,1978II	1.38	1.7
	PIB	1974I,1975I; -1 in 1975II,1975II	2.58	3.3
	PIG	1975I; -1 in 1975II	4.60	5.4
Greece	PCGNW	1986II	13.64	
	PIB	1973II	12.47	4.3
		1974II .	-11.24	3.8
New Zealand	PCGNW	1975I,1976I,1977I; -1 in 1975II,1976II,1977II		1.2
		1985I,1986I; -1 in 1985II,1986II	-3.37	2.3
Norway	PCGNW	1973I,1974I; -1 in 1973II,1974II	-2.91	2.4
		1975I; -1 in 1975II	23.45	10.9
		1984I,1985I; -1 in 1984II,1985II	-3.54	2.9
		1986I; -1 in 1986II	8.02	3.8
	PIB	1985II	17.72	8.2
Spain	PCGNW	1986II	10.90	4.0
Sweden	PCGNW	1980I,1981I; -1 in 1980II,1981II	-3.19	2.7
		1982I,1983I,1984I; -1 in 1982II,1983II,1984II		4.0
	PIB	1978I,1979I,1980I; -1 in 1978II,1979II,1980II	-2.57	4.8
Switzerland	PCGNW		-3.25	2.5
		1976I,1977I; -1 in 1976II, 1977II	-2.04	2.3
		1978I; -1 in 1978II	5.15	4.1
		1980I,1981I; -1 in 1980II,1981II	3.29	3.6

PCGNW Government non-wage consumption deflator PIH Residential construction deflator

PIB Business investment deflator

PIG Government investment deflator

Table 6 Wage/price block simulations (Belgian model)

- 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1										
	1	2	3	4	Yea 5	rs 6	7	8	9	10
	-						, 			
1. Sustained increase in a all cost components (a)		of								
Labour cost	1.0		1.0			1.0			1.0	
Capital cost Import cost	1.0		1.0	1.0		1.0		1.0	1.0	1.(1.(
Private consumption									2.0	`
deflator	0.4	0.9	1.1	1.2	1.1	1.1	1.1	1.0	1.0	1.0
2. Sustained <u>ex ante</u> shoch of wages (a); costs and										
Wage rate	1.1			2.3		2.8			2.8	
Import prices Private consumption	0 1	0.2	0.3	0.4	0.5	0.5	0.5	0.5	0.5	0.
deflator	0.2	0.5	1.0	1.3	1.7	1.8	1.9	1.9	1.8	1.
3. Sustained import price other costs and prices			a);							
Import prices	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0
Wages rate	0.0	0.3	0.4	0.6	0.7	0.8	0.9	1.0	1.0	1.0
Private consumption deflator	0.2	0.4	0.5	0.7	0.8	0.9	0.9	1.0	1.0	1.0
One-period shock to the costs and prices endoge Import prices		0.0		0.0		0.1		•	0.0	• •
Wages rates	0.0	0.1	0.1	0.2	0.2	0.2	0.2	0.1	0.1	0.
Private consumption deflator	0.1	0.1	0.2	0.2	0.2	0.2	0.2	0.1	0.1	0.3

a)

By 1 per cent of the baseline level.
By 1 percentage point of the baseline level. **b**)

Single-country effects of an increase in government expenditure with fixed nominal interest rates and exchange rates (a) (percentage differences from baseline)

Table 7

	Australia	Austria	Belgium	Denmark	Finland	Greece	Ireland	Australia Austria Belgium Denmark Finland Greece Ireland Netherlands New Zealand Norway Spain Sweden Switzerland	New Zealand	l Norway	Spain	Sweden	Witzerland
Real GDF/GNP													
Year 1	1.1	9.0	0.1	6.0	1.3	8.0	4.0	0.5	9.0	9.0	1.1	6.0	8.0
Year 2	1.1	0.5	0.2	1.0	1.3	8.0	0.4	9.0	9.0	0.5	1.4	1.0	9.0
Year 5	9.0	0.2	0.0	9.0	1.0	0.7	0.4	4.0	9.0	0.3	9.0	8.0	0.3
Real total domestic demand	71												
Year 1	1.3	1.3	н. Н	1.3	1.7	1.0	1.0	1.2	1.0	1.2	1.6	1.2	1.2
	1.4	1.5	1.1	1.4	2.0	1.1	1.0	. H	1.0	1.2	2.1	1 .4	. 5.
Year 5	1.6	1.2	1.1	1.1	2.3	1.4	1.1	1.3	1.1	1.7	1.7	1.5	1.2
Real total private investment													
Year 1	0.7		0.1	1.1	1.3	0.3	0.2	9.0	0.1	0.5	1.4	0.5	0.3
	1.0	0.5	0.2	1.3	1.8	9.0	0.5	0 7	0.1	0.3	3.3	1.6	0.5
Year 5	1.6	0.3	0.3	8.0	2.9	3.5	0.5	0.5	0.3	1.3	1.9	1.7	6.0
Consumption deflator													
Year 1	0.3	0.1	0.0	0.2	0.2	0.2	0.0	0.1	0.0	0.3	0.1	0.1	0.2
	1.2	0.2	0.1	9.0	8.0	0.7	0.0	0.2	0.1	1.0	0.4	4 .0	9.0
Year 5	8.8	1.0	4 .0	2.0	3.0	2.1	0.5	0.5	4.0	2.2	1.4	2.4	1.8
Wage rate													
Year 1	4.0	0.2	0.1	0.5	6.0	6.0	0.0	0.1	0.2	9.0	0.3	0.1	0.4
Year 2	1.4	4.0	0.1	8.0	1.9	1.2	0.1	0.2	0.3	1.4	8.0	0.5	8.0
Year 5	ų, S	1.5	4.0	2.4	4.6	2.9	0.5	8.0	6.0	2.4	1.9	3.6	1.9
Unemployment rate (b)													
Year 1	-0.3	-0.3	0.0-	-0.3	-0.3	0.0	-0.0	-0.1	-0.0	0.0-	4.0-	-0.1	-0.1
Year 2	-0.5	-0.2	0.0	- 0 . 4	4.0-	0.0	-0.1	-0.2	0.0-	-0.0	-0.7	-0.2	-0.1
Year 5	0.1	0.1	-0.1	1.0-	-0.1	0.1	-0.2	-0.2	0.0	0.1	-0.2	-0.2	0.1
Current balance (U.S.\$ billion) (b)													
	4.0-	5.0-	6.0-	-0.2	-0.2	-0.4	-0.1	6.0-	-0.1	-0.3	- 0.8	-0.3	4.0-
iear 2 Year 5	12.8	1.0	-2.0	-0.5	8.0-	0.5	0 . 6 . 6 . 6	-1.2 -2.2	10.1	1.1-	-1.1	-1.1	1.5

An increase in government non-wage expenditures equivalent to 1 per cent of baseline real GNP/GDP. Level deviation from baseline.

€ Q

Table 8

Single-country effects of a percentage point decrease in short-term interest rates with floating exchange rates and fixed real government expenditures (percentage differences from baseline)

Real GDP/GNP													
Year 1	0.0	0.2		0.1	0.0	0.1	0.1	0.1	0.0	0.1	0.1	0.0	0.1
		, C		0.2	0.1	0.2	0.2	0.2	0.1	0.3	0.5	0.5	0.3
Year 5	0.1	6.0	0.5	0.0	0.5	9.0	9.0	0.5	0.2	8.0	1.0	9.0	9.0
Real total domestic demand	nd												
Year 1	0.0-	0.2	0.0	0.1	0.0	0.1	0.1	0.0	0.1	0.0	0.0	0.0	0.1
Year 2	0.0	4.0	7.0	0.5	0.1	0.2	0.5	0.0	0.1	0.2	0.3	0.1	0.2
Year 5	0.1	6.0		-0.2	0.2	8.0	8.0	0.2	0.3	0.8	1.0	4.0	0.7
Real total private investment													
Year 1	0.0	0.5	0.2	0.2	0.0	0.2	0.1	0.1	0.0	0.2	0.5	0.1	0.1
Year 2	H. 0	1.0 8.1	0.7 7.4	0.7	0 1.0	0.4 8.4	0.3 1.3	1.0	0.1 0.3	0.6	3.6	0.3 1.2	0 .1 .0 .4
Consumption deflator													
,	•	,	•	ć	6	6	•	ć	9	•	•	,	6
Year I	1.0	T. 6.	4 4	0 0	0.0	0.0	1. E.	0 0 . N	0.0	9.0	. .	0 7 4	4.
	1.3	1.8	1.6	1.3	0.1	0.7	6.0	2.0	4.0-	9. 6.	2.1	1.5	1.9
Wage rate													
Year 1	0.1	0.0	0.1	0.1	0.0	0.1	0.0	0.1	0.0	0.2	0.1	0.1	0.1
Year 2	0.3	0.5	4.0	4.0	0.1	0.2	0.5	ດ.ຮ	0.0-	8.0	4 .	0.5	7
Year 5	1.2	2.0	1.5	1.3	0.2	1.3	1.0	2.0	9.0-	დ დ.	2.3	1.3	6.1
Unemployment rate (a)													
Year 1	0.0-	-0.1	0.0-	-0.0	0.0-	0.0	0.0	0.01	0.0	0.0	0.0-	-0.0	0.0
	0.0-	-0.2	-0.0	0.0	0.0	0.0	0.0-	0.0-	0.1	0.0	-0.1	-0.0	0.0
Year 5	0,1	-0.2	-0.3	6.0	0.0	0.2	-0.1	-0.1	0.1	0.5	0.0-	0.0	0.5
Current balance (U.S.\$ billion) (a)							,						
Year 1	0.0	-0.1	0.0	-0.0	0.0	-0.0	0.0-	-0.1	0.0-	0.0	0.0	0.0-	-0.0
Year 2	0.2	-0.1	0.1	0.0-	0.0	0.0	0.0-	0.0	0.0-	0.0	0	0.0	-0.1
Year 5	6.0	₹.0-	9.0	0.2	-0.1	-0.1	0.0-	4.0	-0.1	0,0	0.0-	o. 0	9.0-

Level deviation from baseline. For New Zealand, the exchange rate is exogenous

Table 9

Single-country effects of a 10 per cent effective depreciation with fixed nominal interest rates and fixed real government expenditures (percentage differences from baseline)

Real GDP/GNP													
Year 1	4.0	2.2		9.0	1.8	0.3	0.7	6.0	0.5	1.1	6.0	0.3	1.3
Year 2	0.9	3.1	2.0	1.0	1.4	0.1	1.6	1.8	4.0	1.6	5.9	1.2	2.2
Year 5	-0.2	0.0		-1.3	1.2	0.0	1.4	4.0	-0.2	0.7	0.1	9.0	0.4
Real total domestic demand	pur												
Year 1	-0.2	0.7	-0.2	0.1	.0.0	-0.7	4.0-	-0.4	-0.8	-0.3	-0.5	-0.2	0.0
Year 2	-0.1	1.2	-0.1	0.4	6.0	-1.2	4 .0	-0.2	-1.2	0.1	1.0	0.4	9.0
Year 5	-0.1	1.0		-2.4	2.8	-0.7	1.4	0.3	-0.5	2.4	8 .0	9.0	1.3
Real total private investment													
Year 1	-0.2	3.5		1.0	0.5	-1.5	-0.3	-0.1	-0.3	0.4	0.5	0.5	-0.1
Year 2	0.5	4.1	2.9	2.4	2.0	-1.7	6.0	1.9	-0.5	1.9	8.	2.5	1.0
Year 5	1.7	1.5		1.1	7.7	5.6	4.0	2.4	9.0	6.9	5.0	3.3	4.5
Consumption deflator													
Year 1	1.7	1.9		2.5	3.0	2.5	1.9	3.0	2.4	2.3	2.0	2.6	2.0
	4.1	4.8	4.9	5.5	5.4	3.7	3.8	6.0	5.5	6.8	4.2	4.1	4.2
Year 5	11.7	15.4	•	7.6	13.9	7.5	8.5	11.3	11.0	14.8	11.7	11.1	12.8
Wage rate													
Year 1	1.4	6.0		1.6	3.2	1.4	1.1	2.1	1.3	2.3	8.0	2.3	1.5
Year 2	4.1	4.8	5.4	4.7	6.4	5.8	3.8	5.5	5.6	7.7	3.3	5.9	* .
Year 5	12.3	17.5	•	9.6	16.7	7.4	4.6	11.8	10.3	13.8	11.7	6.6	13.2
Unemployment rate (a)													
Year 1	-0.3	-1.2	-0.2	9.0-	6.0-	0.0-	-0.1	-0.1	-0.s	-0.1	-0.5	-0.1	-0.3
	9.0-	-1.6	-0.5	9.0-	6.0-	0.0-	-0.3	- 0 . 4	-0.3	-0.1	-1.5	-0.1	-0.4
Year 5	9.0	1.6		1.1	0.3	0.2	4.0-	-0.4	9.0	9.0	8.0	0.1	0.7
(U.S.\$ billion) (a)													
	0.7	-0.2	0.5	10.	4.0	0.1	0.1	-0.1	-0.1	0.3	6.0	0.1	4.0
Year 2	1.8	4.0	•	0.3	0.1	7 .0	0.5	1.5	0.5	9.0	2.3	0.7	6.0
Year 5	9	-0.2	•	1.3	-0.7	0.3	o. 9	2.0	0.5	-0.7	0.0	e.0	-1.4

Level deviation from baseline.

a)

Table 10

Comparison of simulation properties between large and small countries (a)

Year		2	က	7	5	9	7	80	6	10
Fiscal shock										
Real GDP Large countries (b) Small countries	0.9	1.1	1.1	1.0	1.0	0.0 4.0	9.0	0.4	0.3	0.1
rivate consumption deliator Large countries Small countries	0.3	0.9	1.5	2.1	2.5	2.9	3.1	3.1	2.9	2.7
Financial shock										
Real GDP Large countries Small countries	0.1	0.4	0.6	0.7	0.9	1.1	1.1	1.1	1.2	1.1
Private consumption deflator Large countries Small countries	0.1	0.5	1.1	1.7	2.4	3.1	3.9	4.7	3.0	6.4 3.4
Exchange rate shock										
Real GDP Large countries Small countries	0.4 0.9	0.6	0.7	0.7	0.8	0.8	0.5	0.3	0.0	-0.2
Private consumption deflator Large countries Small countries	2.1	4.8 8.8	6.8	8.0	8.8 11.6	9.6	10.2	10.6 10.7	10.5 9.6	10.2

a) Shocks correspond to those in Tables 7 to 9.b) Seven large OECD economies.

ECONOMIC AND STATISTICS DEPARTMENT

WORKING PAPERS

In April 1983, the Economics and Statistics Department initiated a new series of economic studies entitled ESD Working Papers.

The following titles have been circulated:

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