

8

Analyses with Plausible Values

Introduction.....	120
Univariate statistics on plausible values.....	120
The standard error on percentages with PVs.....	123
The standard error on regression coefficients with PVs.....	123
The standard error on correlation coefficients with PVs.....	126
Correlation between two sets of plausible values.....	126
A fatal error shortcut.....	130
An unbiased shortcut.....	131
Conclusion.....	133



INTRODUCTION

As described in Chapters 5 and 6, the cognitive data in PISA are scaled with the Rasch Model and the performance of students is denoted with plausible values (PVs). For minor domains, only one scale is included in the international databases. For major domains,¹ a combined scale and several subscales are provided. For each scale and subscale, five plausible values per student are included in the international databases. This chapter describes how to perform analyses with plausible values.

Since plausible values are mainly used for reporting student performance on cognitive tests, this chapter is mainly useful when conducting analyses on achievement data and their relationships with student or school characteristics.

UNIVARIATE STATISTICS ON PLAUSIBLE VALUES

The computation of a statistic with plausible values always consists of six steps, regardless of the required statistic.

1. The required statistic and its respective standard error have to be computed for each plausible value. In Chapter 7, it was mentioned that 81 estimates were necessary to get the final estimate and its standard error. Therefore, any analysis that involves five plausible values will require 405 estimates. If a mean needs to be estimated, then 405 means will be calculated. The means estimated with the final weight are denoted $\hat{\mu}_1$, $\hat{\mu}_2$, $\hat{\mu}_3$, $\hat{\mu}_4$ and $\hat{\mu}_5$. From the 80 replicates applied on each of the five plausible values, five sampling variances are estimated, denoted respectively $\sigma_{(\hat{\mu}_1)}^2$, $\sigma_{(\hat{\mu}_2)}^2$, $\sigma_{(\hat{\mu}_3)}^2$, $\sigma_{(\hat{\mu}_4)}^2$ and $\sigma_{(\hat{\mu}_5)}^2$. These five mean estimates and their respective sampling variances are provided in Table 8.1.

Table 8.1
The 405 mean estimates

Weight	PV1	PV2	PV3	PV4	PV5
Final	$\hat{\mu}_1$	$\hat{\mu}_2$	$\hat{\mu}_3$	$\hat{\mu}_4$	$\hat{\mu}_5$
Replicate 1	$\hat{\mu}_{1,1}$	$\hat{\mu}_{2,1}$	$\hat{\mu}_{3,1}$	$\hat{\mu}_{4,1}$	$\hat{\mu}_{5,1}$
Replicate 2	$\hat{\mu}_{1,2}$	$\hat{\mu}_{2,2}$	$\hat{\mu}_{3,2}$	$\hat{\mu}_{4,2}$	$\hat{\mu}_{5,2}$
Replicate 3	$\hat{\mu}_{1,3}$	$\hat{\mu}_{2,3}$	$\hat{\mu}_{3,3}$	$\hat{\mu}_{4,3}$	$\hat{\mu}_{5,3}$
.....
.....
Replicate 80	$\hat{\mu}_{1,80}$	$\hat{\mu}_{2,80}$	$\hat{\mu}_{3,80}$	$\hat{\mu}_{4,80}$	$\hat{\mu}_{5,80}$
Sampling variance	$\sigma_{(\hat{\mu}_1)}^2$	$\sigma_{(\hat{\mu}_2)}^2$	$\sigma_{(\hat{\mu}_3)}^2$	$\sigma_{(\hat{\mu}_4)}^2$	$\sigma_{(\hat{\mu}_5)}^2$

2. The final mean estimate is equal to the average of the five mean estimates, *i.e.* $\hat{\mu} = \frac{1}{5}(\hat{\mu}_1 + \hat{\mu}_2 + \hat{\mu}_3 + \hat{\mu}_4 + \hat{\mu}_5)$.

3. The final sampling variance is equal to the average of the five sampling variances,

$$i.e. \sigma_{(\hat{\mu})}^2 = \frac{1}{5}(\sigma_{(\hat{\mu}_1)}^2 + \sigma_{(\hat{\mu}_2)}^2 + \sigma_{(\hat{\mu}_3)}^2 + \sigma_{(\hat{\mu}_4)}^2 + \sigma_{(\hat{\mu}_5)}^2)$$

4. The imputation variance, also denoted measurement error variance, is computed as $\sigma_{(test)}^2 = \frac{1}{4} \sum_{i=1}^5 (\hat{\mu}_i - \hat{\mu})^2$. Indeed, as PISA returns five plausible values per scale, then $\sigma_{(test)}^2 = \frac{1}{M-1} \sum_{i=1}^M (\hat{\mu}_i - \hat{\mu})^2 = \frac{1}{4} \sum_{i=1}^5 (\hat{\mu}_i - \hat{\mu})^2$ with M being the number of plausible values. This formula is similar to the one used for the estimation of a population variance, except that in this particular case, observations are not compared with the population mean, but each PV mean is compared with the final mean estimate.



5. The sampling variance and the imputation variance are combined to obtain the final error variance as

$$\sigma_{(error)}^2 = \sigma_{(\hat{\mu})}^2 + (1.2\sigma_{(test)}^2)$$

$$\text{Indeed, } \sigma_{(error)}^2 = \sigma_{(\hat{\mu})}^2 + \left(\left(1 + \frac{1}{M} \right) \sigma_{(test)}^2 \right) = \sigma_{(\hat{\mu})}^2 + \left(\left(1 + \frac{1}{5} \right) \sigma_{(test)}^2 \right) = \sigma_{(\hat{\mu})}^2 + ((1.2)\sigma_{(test)}^2)$$

6. The standard error is equal to the square root of the error variance.

The mean estimate on the science scale and its respective standard error for the PISA 2006 Belgian data can be computed. The macro described in Chapter 7 and labelled PROC_MEANS_NO_PV can be sequentially used five times, and the results can be combined in an Excel® spreadsheet. Table 8.2 presents the different PV means, their respective sampling variances, as well as the mean estimates on the first and last replicates.

Table 8.2

Mean estimates and their respective sampling variances on the science scale for Belgium (PISA 2006)

Weight	PV1	PV2	PV3	PV4	PV5
Final	510.18	510.58	510.36	510.62	510.09
Replicate 1	509.06	509.63	509.64	509.72	509.03
.....
Replicate 80	511.38	512.03	511.63	512	511.28
Sampling variance	(2.47) ²	(2.42) ²	(2.50) ²	(2.45) ²	(2.52) ²

Box 8.1 presents the SAS® syntax for running sequentially the PROC_MEANS_NO_PV macro described in Chapter 7.

Box 8.1 SAS® syntax for computing the mean on the science scale by using the PROC_MEANS_NO_PV macro (e.g. PISA 2006)

```
libname PISA2003 "c:\pisa\2003\data\";
libname PISA2006 "c:\pisa\2006\data\";
options nofmterr notes;
run;
data temp1;
    set pisa2006.stu;
    if (cnt in ("BEL"));
    w_fstr0=w_fstuw;
    science1=pv1scie;
    science2=pv2scie;
    science3=pv3scie;
    science4=pv4scie;
    science5=pv5scie;
    if (st04q01=1) then gender=1;
    if (st04q01=2) then gender=0;
    keep cnt schoolid stidstd w_fstr0-w_fstr80 science1-science5
        st04q01 gender hisei bsmj;
run;
%include "c:\pisa\macro\proc_means_no_pv.sas";
%macro repeat;
%do kk=1 %to 5;
%BRR_PROCMEAN(INFILE=temp1,
    REPLI_ROOT =w_fstr,
    BYVAR=cnt,
    VAR=science&kk,
    STAT=mean,
    LIMIT=yes,
    LIMIT_CRITERIA=100 10 5 1,
    ID_SCHOOL=schoolid,
    OUTFILE =exercise&kk);
run;
%end;
%mend;
%repeat;
run;
```



With the results in the exercise1 to exercise5 files, the final mean estimate for Belgium on the combined science scale is computed as:

$$\hat{\mu} = \frac{1}{5}(\hat{\mu}_1 + \hat{\mu}_2 + \hat{\mu}_3 + \hat{\mu}_4 + \hat{\mu}_5), \text{ i.e. } \hat{\mu} = \frac{(510.18 + 510.58 + 510.36 + 510.62 + 510.09)}{5} = 510.4$$

The final sampling variance on the mean estimate for the combined science literacy scale is equal to:

$$\sigma_{(\hat{\mu})}^2 = \frac{1}{5}(\sigma_{(\hat{\mu}_1)}^2 + \sigma_{(\hat{\mu}_2)}^2 + \sigma_{(\hat{\mu}_3)}^2 + \sigma_{(\hat{\mu}_4)}^2 + \sigma_{(\hat{\mu}_5)}^2), \text{ i.e. } \sigma_{(\hat{\mu})}^2 = \frac{(2.47)^2 + (2.42)^2 + (2.50)^2 + (2.45)^2 + (2.52)^2}{5} = 6.11$$

The imputation variance is equal to:

$$\sigma_{(test)}^2 = \frac{1}{4} \sum_{i=1}^5 (\hat{\mu}_i - \hat{\mu})^2, \text{ i.e. } \sigma_{(test)}^2 = \frac{[(510.18 - 510.4)^2 + (510.58 - 510.4)^2 + \dots + (510.09 - 510.4)^2]}{4} = 0.055$$

The final error variance is equal to:

$$\sigma_{(error)}^2 = \sigma_{(\hat{\mu})}^2 + (1.2\sigma_{(test)}^2), \text{ i.e. } \sigma_{(error)}^2 = 6.11 + (1.2 * 0.055) = 6.17$$

The final standard error is therefore equal to:

$$SE = \sqrt{\sigma_{(error)}^2} = \sqrt{6.17} = 2.48$$

Sequentially running the PROC_MEANS_NO_PV macro five times and combining the results can be avoided: a SAS® macro has been developed for dealing with plausible values (see Box 8.2). This macro also computes the:

- five mean estimates,
- five sampling variances,
- imputation variance,
- final standard error by combining the final sampling variance and the imputation variance.

Box 8.2 SAS® syntax for computing the mean and its standard error on PVs (e.g. PISA 2006)

```
%include "c:\pisa\macro\proc_means_pv.sas";
%BRR_PROCMEAN_PV(INFILE=templ,
                  REPLI_ROOT =w_fstr,
                  BYVAR=cnt,
                  PV_ROOT=science,
                  STAT=mean,
                  LIMIT=yes,
                  LIMIT_CRITERIA=100 10 5 1,
                  ID_SCHOOL=schoolid,
                  OUTFILE =exercise6);
run;
```

Besides the seven arguments common to all SAS® macros described in this manual, *i.e.* (i) INFILE=; (ii) REPLI_ROOT=; (iii) BYVAR=; (iv) OUTFILE=; (v) LIMIT=; (vi) LIMIT_CRITERIA=; and (vii) ID_SCHOOL=, two additional ones need to be defined:

- The root of the variable names for the five plausible values. In the PISA database, PV names are usually PV1READ, PV2READ, ..., PV1MATH, PV2MATH, ..., PV1SCIE, PV2SCIE, These variable names cannot be used directly by the macro, as it will automatically add the numbers 1 to 5 at the end of the root variable name. Therefore, in the "data" statement in Box 8.1, new variables are created to fit the macro requirements. When calling the macro, the argument PV_ROOT will be equal to SCIENCE.
- The STAT argument to specify the requested statistics. Available statistics have been described in Chapter 7.



The structure of the output data file exercise6 is presented in Table 8.3.

Table 8.3
Output data file exercise6 from Box 8.2

CNT	STAT	SESTAT	FLAG_STUD	FLAG_SCH	FLAG_PCT
BEL	510.4	2.48	0	0	0

Similar to the SAS[®] macros described in the previous chapter, more than one breakdown variable can be used. For instance, if one wants to determine whether the dispersion of the science performance is larger for females than for males, the macro BRR_PROCMEAN_PV can be used as shown in Box 8.3.

Box 8.3 SAS[®] syntax for computing the standard deviation and its standard error on PVs by gender (e.g. PISA 2006)

```
%BRR_PROCMEAN_PV ( INFILE=temp1,
                    REPLI_ROOT=w_fstr,
                    BYVAR=cnt st04q01,
                    PV_ROOT=science,
                    STAT=std,
                    LIMIT=yes,
                    LIMIT_CRITERIA=100 10 5 1,
                    ID_SCHOOL=schoolid,
                    OUTFILE=exercise7) ;

run;
```

The structure of the output data file is presented in Table 8.4.

Table 8.4
Output data file exercise7 from Box 8.3

CNT	ST04Q01	STAT	SESTAT	FLAG_STUD	FLAG_SCH	FLAG_PCT
BEL	1	96.1	2.23	0	0	0
BEL	2	102.9	2.50	0	0	0

According to Table 8.4, the standard deviation (STAT) is larger for males than for females. Unfortunately, the information of the two standard errors in SESTAT is not enough to conduct a test of the equality for these two standard deviation coefficients since the standard deviation estimates for males and females may be correlated. The detail of the standard error on difference will be presented in Chapter 11.

THE STANDARD ERROR ON PERCENTAGES WITH PVs

The PROC_FREQ_NO_PV first presented in Chapter 7 was developed for the computation of percentages and their respective standard errors. Chapter 9 will deal with the application of this macro to plausible values: an entire chapter needs to be devoted to this type of analysis because of the issues involved.

THE STANDARD ERROR ON REGRESSION COEFFICIENTS WITH PVs

Suppose that the statistical effect of gender and student socio-economic background on the performance in science needs to be estimated. Just like estimating a mean, this question can be solved by sequentially applying five times the PROC_REG_NO_PV macro described in Chapter 7. Box 8.4 presents the SAS[®] syntax for such an approach.



Box 8.4 SAS® syntax for computing regression coefficients and their standard errors on PVs by using the PROC_REG_NO_PV macro (e.g. PISA 2006)

```
%include "c:\pisa\macro\proc_reg_no_pv.sas";
%macro repeat;
%do kk=1 %to 5;
%BRR_REG( INFILE=templ,
          REPLI_ROOT=w_fstr,
          BYVAR=cnt,
          VARDEP=science&kk,
          EXPLICA=hisei gender,
          LIMIT=yes,
          LIMIT_CRITERIA=100 10 5 1,
          ID_SCHOOL=schoolid,
          OUTFILE=exercise&kk);
run;
%end;
%mend;
%repeat;
run;
```

Just like the computation of a mean and its standard error, the computation of regression coefficients and their respective standard errors consists of six steps:

1. For each plausible value and for each explanatory variable, regression coefficients are computed with the final and the 80 replicate weights. Thus, 405 regression coefficients per explanatory variable will be computed. The PROC_REG_NO_PV macro applied sequentially five times will return, per explanatory variable, five estimates, denoted, $\hat{\beta}_1, \dots, \hat{\beta}_5$ and five standard errors, denoted $\sigma_{(\hat{\beta}_1)}, \dots, \sigma_{(\hat{\beta}_5)}$. Table 8.5 presents the mathematical expression for these 405 estimates and Table 8.6 presents some of the values for the 405 regression coefficients obtained on the Belgian data for the HISEI (international socio-economic index of occupational status) variable.

2. The final regression coefficient estimate is equal to $\hat{\beta} = \frac{\hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3 + \hat{\beta}_4 + \hat{\beta}_5}{5}$

$$\text{i.e. for HISEI } \hat{\beta} = \frac{2.22 + 2.24 + 2.25 + 2.24 + 2.26}{5} = 2.24$$

3. The final sampling variance estimate is equal to $\sigma_{(\hat{\beta})}^2 = \frac{1}{5} (\sigma_{(\hat{\beta}_1)}^2 + \sigma_{(\hat{\beta}_2)}^2 + \sigma_{(\hat{\beta}_3)}^2 + \sigma_{(\hat{\beta}_4)}^2 + \sigma_{(\hat{\beta}_5)}^2)$

$$\text{i.e. for HISEI } \sigma_{(\hat{\beta})}^2 = \frac{(0.09)^2 + (0.09)^2 + (0.09)^2 + (0.09)^2 + (0.09)^2}{5} = 0.0081$$

4. The imputation variance is equal to $\sigma_{(test)}^2 = \frac{1}{4} \sum_{i=1}^5 (\hat{\beta}_i - \hat{\beta})^2$

$$\text{i.e. for HISEI } \sigma_{(test)}^2 = \frac{(2.22 - 2.24)^2 + (2.24 - 2.24)^2 + \dots + (2.26 - 2.24)^2}{4} = 0.0002$$

5. The final error variance is equal to $\sigma_{(error)}^2 = \sigma_{(\hat{\beta})}^2 + (1.2 \sigma_{(test)}^2)$

$$\text{i.e. for HISEI } \sigma_{(error)}^2 = 0.0081 + (1.2 * 0.0002) = 0.0084$$

6. The final standard error is equal to $SE = \sqrt{\sigma_{(error)}^2} = \sqrt{0.0084} = 0.10$

As 2.24 divided by 0.10 is about 22.4, the regression coefficient for HISEI is significantly different from 0.



Table 8.5
The 450 regression coefficient estimates

Weight	PV1	PV2	PV3	PV4	PV5
Final	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$	$\hat{\beta}_5$
Replicate 1	$\hat{\beta}_{1,1}$	$\hat{\beta}_{2,1}$	$\hat{\beta}_{3,1}$	$\hat{\beta}_{4,1}$	$\hat{\beta}_{5,1}$
Replicate 2	$\hat{\beta}_{1,2}$	$\hat{\beta}_{2,2}$	$\hat{\beta}_{3,2}$	$\hat{\beta}_{4,2}$	$\hat{\beta}_{5,2}$
Replicate 3	$\hat{\beta}_{1,3}$	$\hat{\beta}_{2,3}$	$\hat{\beta}_{3,3}$	$\hat{\beta}_{4,3}$	$\hat{\beta}_{5,3}$
.....
.....
Replicate 80	$\hat{\beta}_{1,80}$	$\hat{\beta}_{2,80}$	$\hat{\beta}_{3,80}$	$\hat{\beta}_{4,80}$	$\hat{\beta}_{5,80}$
Sampling variance	$\sigma^2_{(\hat{\beta}_1)}$	$\sigma^2_{(\hat{\beta}_2)}$	$\sigma^2_{(\hat{\beta}_3)}$	$\sigma^2_{(\hat{\beta}_4)}$	$\sigma^2_{(\hat{\beta}_5)}$

Table 8.6
HISEI regression coefficient estimates and their respective sampling variance on the science scale in Belgium after accounting for gender (PISA 2006)

Weight	PV1	PV2	PV3	PV4	PV5
Final	2.22	2.24	2.25	2.24	2.26
Replicate 1	2.22	2.22	2.24	2.23	2.27
.....
Replicate 80	2.14	2.16	2.18	2.18	2.18
Sampling variance	(0.09) ²	(0.09) ²	(0.09) ²	(0.09) ²	(0.09) ²

A SAS[®] macro PROC_REG_PV has also been developed for regression analyses with plausible values as dependent variables. The SAS[®] syntax is presented in Box 8.5.

Box 8.5 **SAS[®] syntax for running the simple linear regression macro with PVs (e.g. PISA 2006)**

```
%include "c:\pisa\macro\proc_reg_pv.sas";
%BRR_REG_PV(INFILE=templ,
             REPLI_ROOT=w_fstr,
             EXPLICA=hisei gender,
             BYVAR=cnt,
             PV_ROOT=science,
             LIMIT=yes,
             LIMIT_CRITERIA=100 10 5 1,
             ID_SCHOOL=schoolid,
             OUTFILE=exercise8);
run;
```

Besides the arguments common to all macros, the root of the plausible value variable names has to be specified, in addition to the list of independent variables. The structure of the output data file is presented in Table 8.7.

Table 8.7
Output data file exercise8 from Box 8.5

CNT	CLASS	STAT	SESTAT
BEL	Intercept	403.39	5.74
BEL	HISEI	2.24	0.10
BEL	Gender	0.10	3.24
BEL	_RSQ_	0.15	0.01



A quick overview of these results shows that the intercept and the regression coefficient for HISEI are significantly different from 0. However, the regression coefficient for GENDER does not statistically differ from 0.

Similar to the PROC_REG_NO_PV macro, the PROC_REG_PV macro also returns the outcomes of the sampling size requirement analysis in a separate file. The file name consists of: (i) the name of the output files that contains the regression parameters; and (ii) the name of the output file followed with _CRITERIA.

THE STANDARD ERROR ON CORRELATION COEFFICIENTS WITH PVS

A SAS® macro has also been developed to compute the correlation between a set of plausible values and another variable. The SAS syntax for running this macro is presented in Box 8.6 and the structure of the output data file is presented in Table 8.8.

Box 8.6 SAS® syntax for running the correlation macro with PVs (e.g. PISA 2006)

```
%include "c:\pisa\macro\proc_corr_pv.sas";

%BRR_CORR_PV( INFILE=temp1,
              REPLI_ROOT=w_fstr,
              BYVAR=cnt ,
              EXPLICA=hisei,
              PV_ROOT=science,
              LIMIT=yes,
              LIMIT_CRITERIA=100 10 5 1,
              ID_SCHOOL=schoolid,
              OUTFILE=exercise9);

run;
```

Table 8.8

Output data file exercise9 from Box 8.6

CNT	STAT	SESTAT	FLAG_STUD	FLAG_SCH	FLAG_PCT
BEL	0.38	0.02	0	0	0

CORRELATION BETWEEN TWO SETS OF PLAUSIBLE VALUES

Some researchers may be interested in the correlation between the different PISA domains and subdomains. For instance, some might want to compute the correlation between the reading subdomains or between the mathematics subdomains, or between reading and mathematics using the PISA 2000, PISA 2003 and PISA 2006 databases.

As described in Chapter 5, the PISA assessment used incomplete assessment designs, *i.e.* students are required to answer a subset of the item battery. Further, while all students were assessed in the major domain, only a subset of students was assessed in minor domains.

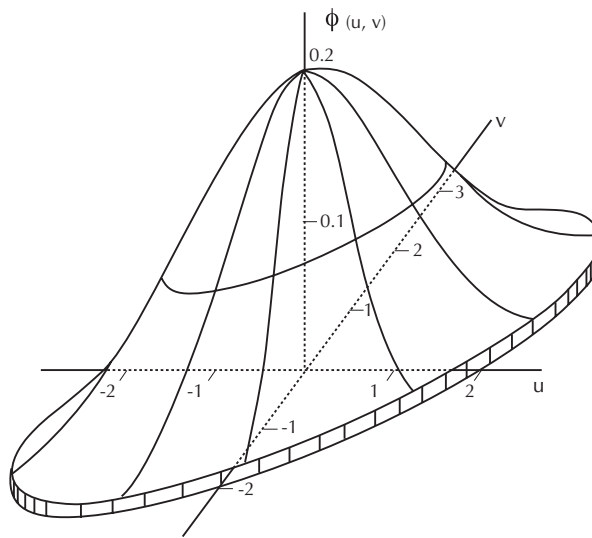
PISA 2000 included plausible values of a minor domain only for students who answered questions for that minor domain. Therefore, using the PISA 2000 database to compute the correlation between reading and mathematics, for example, would require working on a subset of students.²

To facilitate secondary analyses, PISA 2003 returned plausible values for all domains and for all students, regardless of whether they were actually assessed or not. Ignoring the assessment status is possible, because the cognitive data in PISA are scaled according to multi-dimensional models.



Since this is easier to illustrate graphically, let's suppose that only two domains were assessed, *i.e.* mathematics/quantity and mathematics/space and shape. If the mathematics/quantity and mathematics/space and shape materials were scaled independently, the correlation between the two subdomains would be largely underestimated. In order to avoid this problem, both materials are scaled together. The model will build a two-dimensional posterior distribution, instead of two one-dimensional posterior distributions as described in Chapter 6. Figure 8.1 graphically presents a two-dimensional normal distribution.

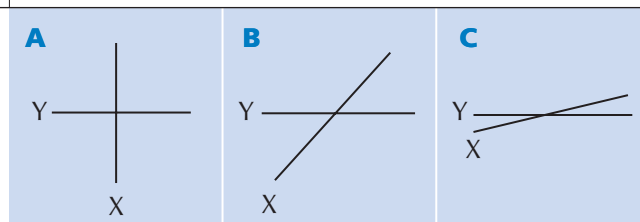
Figure 8.1
A two-dimensional distribution



To correctly describe such distributions, two means, two variances, and one correlation are needed. If the correlation is equal to 0, then the two axes will be orthogonal. As the absolute value of the correlation starts to increase, the angle formed by the two axes becomes less than 90 degrees (see Figure 8.2, part A).³ Two axes perfectly overlapping would represent a correlation of 1.0 (or -1.0). These different cases are illustrated in Figure 8.2.

With a two-dimensional model, the first plausible value for mathematics/quantity will be drawn at the same time as the first plausible value for mathematics/space and shape. Per student, this will consist of randomly drawing one dot in the scatter plot. The values of the two plausible values will be the coordinates of the dot on the two axes. The same procedure is applied for the second, third, fourth, and fifth plausible values.

Figure 8.2
Axes for two-dimensional normal distributions





As the PISA domains and subdomains highly correlate, as shown by the graph on the far right in Figure 8.2, it is very unlikely for a student to get a high score for the first plausible value in mathematics/quantity (PV1MATH4) and a low score for the first plausible value in mathematics/space and shape (PV1MATH1). If plausible values were drawn independently for these two mathematics subdomains, such a case would be possible and therefore the correlation would be underestimated.

Since each draw is independent, to calculate the correlation between the two domains, the correlation between each set of plausible value, *i.e.* correlation between two plausible values with the same number, needs to be computed. In the PISA 2003 example of correlation between PV1MATH1 to PV5MATH1 and PV1MATH4 to PV5MATH4, the following values need to be computed:

- PV1MATH1 and PV1MATH4,
- PV2MATH1 and PV2MATH4,
- PV3MATH1 and PV3MATH4,
- PV4MATH1 and PV4MATH4,
- PV5MATH1 and PV5 MATH4.

Table 8.9 presents all possible 25 correlation coefficients between the 5 plausible values in mathematics/quantity (MATH4) and mathematics/space and shape (MATH1), respectively, for Germany in PISA 2003. As expected, the correlation coefficients on the diagonal of the square matrix are substantially higher than the other correlation coefficients. The final correlation estimate between these two mathematics subdomains will be the average of the five correlation coefficients on the diagonal.

Table 8.9

Correlation between the five plausible values for each domain, mathematics/quantity and mathematics/space and shape

	PV1MATH1	PV2MATH1	PV3MATH1	PV4MATH1	PV5MATH1
PV1MATH4	0.90	0.83	0.84	0.84	0.83
PV2MATH4	0.83	0.90	0.84	0.84	0.83
PV3MATH4	0.84	0.83	0.90	0.84	0.83
PV4MATH4	0.83	0.83	0.84	0.90	0.83
PV5MATH4	0.83	0.83	0.84	0.84	0.90

The standard error on this correlation estimate can be easily obtained by applying five times sequentially the PROC_CORR_NO_PV macro described in Chapter 7. The SAS[®] syntax is given in Box 8.7 and five correlation coefficients and respective sampling variances, produced by the syntax in Box 8.7 are presented in Table 8.10.

The final correlation estimate is equal to:

$$\hat{\rho} = \frac{\hat{\rho}_1 + \hat{\rho}_2 + \hat{\rho}_3 + \hat{\rho}_4 + \hat{\rho}_5}{5}, \text{ i.e. } \hat{\rho} = \frac{0.8953 + 0.8964 + \dots + 0.8958}{5} = 0.8970$$

The final sampling variance is equal to:

$$\sigma_{(\hat{\rho})}^2 = \frac{\sum_{i=1}^5 \sigma_{(\hat{\rho}_i)}^2}{5}, \text{ i.e. } \sigma_{(\hat{\rho})}^2 = \frac{(0.0040)^2 + (0.0033)^2 + \dots + (0.0038)^2}{5} = 0.000013$$



The measurement variance can be estimated as:

$$\sigma_{(test)}^2 = \frac{1}{4} \sum_{i=1}^5 (\hat{\rho}_i - \hat{\rho})^2 = 0.000003$$

The final variance is equal to:

$$\sigma_{(error)}^2 = \sigma_{(\hat{\rho})}^2 + (1.2 \sigma_{(test)}^2) = 0.000017$$

The final standard error is equal to:

$$SE = \sqrt{\sigma_{(error)}^2} = \sqrt{0.000017} = 0.004116$$

Box 8.7 **SAS® syntax for the computation of the correlation between mathematics/quantity and mathematics/space and shape by using the PROC_CORR_NO_PV macro^a (e.g. PISA 2003)**

```
data temp2;
    set pisa2003.stud;
    if (cnt="DEU") ;
    w_fstr0=w_fstuw;
    keep  cnt schoolid stidstd w_fstr0-w_fstr80
         pv1math1 pv2math1 pv3math1 pv4math1 pv5math1
         pv1math4 pv2math4 pv3math4 pv4math4 pv5math4;
run;
%include "c:\pisa\macro\proc_corr_no_pv.sas";
%macro repeat;
%do kk=1 %to 5;
%BRR_CORR(
    INFILE=temp2,
    REPLI_ROOT=w_fstr,
    BYVAR=cnt,
    VAR1=pv&kk.math1,
    VAR2=pv&kk.math4,
    LIMIT=no,
    LIMIT_CRITERIA=,
    ID_SCHOOL=,
    OUTFILE=exercise&kk);
run;
%end;
%mend;
%repeat;
run;
```

a. In the PROC_CORR_NO_PV macro, the statement of FORMAT STAT F5.2 should be changed to FORMAT STAT F8.4 to obtain the correlation coefficients with four decimal points.

Table 8.10

The five correlation estimates between mathematics/quantity and mathematics/space and shape and their respective sampling variance

	PV1	PV2	PV3	PV4	PV5
Correlation	0.8953	0.8964	0.8996	0.8978	0.8958
Sampling variance	(0.0040) ²	(0.0033) ²	(0.0034) ²	(0.0037) ²	(0.0038) ²



The computation of the correlation between two domains or between a subdomain and a domain might be problematic in some cases in the PISA databases. PISA 2000 used two scaling models:

- a three-dimensional model with mathematics, reading and science;
- a five-dimensional model with mathematics, reading/retrieving information, reading/interpreting texts, reading/reflection and evaluation, and science.

PISA 2003 also used two scaling models:

- a four-dimensional model with mathematics, problem solving, reading and science;
- a seven-dimensional model with mathematics/space and shape, mathematics/change and relationships, mathematics/uncertainty, mathematics/quantity, problem solving, reading and science.

PISA 2006 used two scaling models as well:

- a five-dimensional model with mathematics, reading, science and the two attitudinal scales;
- a five-dimensional model with mathematics, reading, and the three science scales (the identifying scientific issues scale, the explaining phenomena scientifically scale, and the using scientific evidence scale).

The PISA databases should contain two sets of plausible values for each of the minor domains. As this would be too confusing, only one set was provided. Therefore, the correlation coefficients are underestimated.

This can be confirmed by examining the data. In the case of a minor domain and a subscale of the major domain, the correlation coefficients on the diagonal do not differ from the other correlations, since these two sets of plausible values were generated by two different models.

In PISA 2006, as well as in PISA 2000 and in PISA 2003, the plausible values for the minor domains included in the databases were generated with the major domain as a combined scale. This means that:

- The correlation between a minor domain and the combined scale of the major domain can be computed.
- The correlation between two minor domains can be computed.
- The correlation between the subdomains can be computed.
- It is not possible to compute the correlation between minor domains and one of the subscales of the major domain.

A FATAL ERROR SHORTCUT

A common fatal error when analysing with plausible values involves computing the mean of the five plausible values, before further analysis.

In Chapter 6, the expected *a posteriori* (EAP) student performance estimator was described. As a reminder, the EAP estimator is equal to the mean of the posterior distribution. Therefore, computing the mean of the five plausible values at the student level is more or less equal to the EAP estimate.

In Chapter 6, the efficiency of the EAP estimator was also compared with the weighted likelihood estimate (WLE) and the plausible values for some statistics estimations. It was indicated that the EAP estimator:

- underestimates the standard deviation,
- overestimates the correlation between the student performance and some background variables,
- underestimates the within-school variance.



Therefore, computing the mean of the five plausible values and then computing statistics on this new score would bias the results just as the EAP does. Table 8.11 provides, per country, the standard deviation of the combined mathematics scale in PISA 2003 using the correct method, as described in this chapter, and the incorrect method of averaging the five plausible values at the student level and then computing the standard deviation on this new score. The result of the latter is denoted as pseudo-EAP.

As shown by Table 8.11, the pseudo-EAP underestimates the standard deviation.

Table 8.11

Standard deviations for mathematics scale using the correct method (plausible values) and by averaging the plausible values at the student level (pseudo-EAP) (PISA 2003)

	Plausible values	Pseudo-EAP
AUS	95.42	91.90
AUT	93.09	89.91
BEL	109.88	106.65
CAN	87.11	83.37
CHE	98.38	94.97
CZE	95.94	92.50
DEU	102.59	99.54
DNK	91.32	87.52
ESP	88.47	84.52
FIN	83.68	79.77
FRA	91.70	88.07
GBR	92.26	89.18
GRC	93.83	89.49
HUN	93.51	89.71
IRL	85.26	82.03
ISL	90.36	86.55
ITA	95.69	92.00
JPN	100.54	96.96
KOR	92.38	89.07
LUX	91.86	88.28
MEX	85.44	80.52
NLD	92.52	89.89
NOR	92.04	88.31
NZL	98.29	95.07
POL	90.24	86.49
PRT	87.63	83.91
SVK	93.31	89.86
SWE	94.75	91.07
TUR	104.74	100.79
USA	95.25	92.12

The analysis process should always aggregate the results of the five plausible values at the latest stage, *i.e.* the statistic that has to be reported is computed five times, then these five statistics are combined.

AN UNBIASED SHORTCUT

Table 8.1 and Table 8.5 respectively give the 405 mean and regression coefficient estimates needed for computing a mean or regression coefficient final estimate and the respective standard errors.

On average, analysing one plausible value instead of five plausible values provides unbiased population estimates as well as unbiased sampling variances on these estimates. It will not be possible to estimate the imputation variance using this method, however.

Therefore, an unbiased shortcut consists of:

- computing, using one of the five plausible values, the statistical estimate and its sampling variance by using the final student weight as well as the 80 replicate weights;
- computing the statistical estimate by using the final student weight on the four other plausible values;



- computing the final statistical estimate by averaging the plausible value statistical estimates;
- computing the imputation variance, as previously described;
- combining the imputation variance and the sampling variance, as previously described.

This unbiased shortcut is presented in Table 8.12 for the estimation of a mean and its standard error. This shortcut only requires the computation of 85 estimates instead of 405. The final estimate of this shortcut will be equal to the one obtained with the long procedure, but the standard error might differ slightly.

Table 8.12
Unbiased shortcut for a population estimate and its standard error

Weight	PV1	PV2	PV3	PV4	PV5
Final	$\hat{\mu}_1$	$\hat{\mu}_2$	$\hat{\mu}_3$	$\hat{\mu}_4$	$\hat{\mu}_5$
Replicate 1	$\mu_{1,1}$				
Replicate 2	$\mu_{1,2}$				
Replicate 3	$\mu_{1,3}$				
.....				
.....				
Replicate 80	$\mu_{1,80}$				
Sampling variance	$\sigma^2_{(\hat{\mu}_1)}$				

Table 8.13
Standard errors from the full and shortcut computation (PISA 2006)

	Mean estimate in science		Immigrant students mean estimate in science		Student flag (50)
	Full computation	Shortcut computation	Full computation	Shortcut computation	
	S.E.	S.E.	S.E.	S.E.	
AUS	2.26	2.21	5.75	5.77	0
AUT	3.92	3.98	10.92	11.00	0
BEL	2.48	2.49	8.34	8.65	0
CAN	2.03	2.00	5.24	5.25	0
CHE	3.16	3.25	6.94	7.13	0
CZE	3.48	3.40	15.89	15.48	0
DEU	3.80	3.80	8.80	8.62	0
DNK	3.11	3.06	8.05	8.06	0
ESP	2.57	2.60	7.23	7.43	0
FIN	2.02	2.05	16.28	15.76	0
FRA	3.36	3.34	10.05	9.94	0
GBR	2.29	2.23	14.74	15.07	0
GRC	3.23	3.29	10.33	10.73	0
HUN	2.68	2.63	14.57	13.15	0
IRL	3.19	3.18	14.56	13.95	0
ISL	1.64	1.59	13.94	13.73	0
ITA	2.02	2.02	8.18	8.30	0
JPN	3.37	3.45	36.57	37.54	1
KOR	3.36	3.41			
LUX	1.05	1.14	3.73	3.82	0
MEX	2.71	2.64	7.61	6.78	0
NLD	2.74	2.77	10.16	10.18	0
NOR	3.11	3.07	11.21	11.34	0
NZL	2.69	2.67	6.58	6.79	0
POL	2.34	2.37	53.97	49.11	1
PRT	3.02	3.02	11.09	10.64	0
SVK	2.59	2.57	47.09	49.10	1
SWE	2.37	2.28	8.09	8.29	0
TUR	3.84	3.82	18.20	18.18	1
USA	4.22	4.20	7.89	8.15	0



Table 8.13 presents the standard errors for the country mean estimate in student performance on the science scale (PISA 2006) as well as the immigrant subpopulations. Two standard errors are provided per estimate: the standard error that results from the strict application of the recommended procedure; and the standard errors using the described shortcut. In all or nearly all cases, the difference between the two standard error estimates is negligible, even based on less than 50 students.

CONCLUSION

This chapter described the different steps for analysing data with plausible values. It also provided some SAS[®] macros to facilitate the computations.

Attention was also drawn to a common error that consists of computing the average of the plausible values at the student level and adding this value to the database to be used as the student score in analyses. The correct method involves the averaging process which should always occur at the latest stage, that is on the statistic that will be reported.

The particular issue of analysing two sets of plausible values was also presented in the case of a correlation. The procedure that was applied can also be extended to other linear or non-linear modelling, such as a linear regression analysis.

Finally, an unbiased shortcut was described, one that is useful for time-consuming procedures, e.g. multilevel procedures.

Notes

1. Reading was the major domain in PISA 2000, mathematics in PISA 2003 and science in PISA 2006.
2. For more information, see the *Manual for the PISA 2000 Database* (OECD, 2002b).
3. A correlation coefficient can be expressed by the cosines of the angle formed by the two variables.



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Table of contents

FOREWORD	3
USER'S GUIDE	17
CHAPTER 1 THE USEFULNESS OF PISA DATA FOR POLICY MAKERS, RESEARCHERS AND EXPERTS ON METHODOLOGY	19
PISA – an overview	20
▪ The PISA surveys.....	20
How can PISA contribute to educational policy, practice and research?	22
▪ Key results from PISA 2000, PISA 2003 and PISA 2006.....	23
Further analyses of PISA datasets	25
▪ Contextual framework of PISA 2006.....	28
▪ Influence of the methodology on outcomes.....	31
CHAPTER 2 EXPLORATORY ANALYSIS PROCEDURES	35
Introduction	36
Weights	36
Replicates for computing the standard error	39
Plausible values	43
Conclusion	46
CHAPTER 3 SAMPLE WEIGHTS	49
Introduction	50
Weights for simple random samples	51
Sampling designs for education surveys	53
Why do the PISA weights vary?	57
Conclusion	58
CHAPTER 4 REPLICATE WEIGHTS	59
Introduction	60
Sampling variance for simple random sampling	60
Sampling variance for two-stage sampling	65
Replication methods for simple random samples	70
Replication methods for two-stage samples	72
▪ The Jackknife for unstratified two-stage sample designs.....	72
▪ The Jackknife for stratified two-stage sample designs.....	73
▪ The Balanced Repeated Replication method.....	74
Other procedures for accounting for clustered samples	76
Conclusion	76



CHAPTER 5 THE RASCH MODEL	79
Introduction	80
How can the information be summarised?	80
The Rasch Model for dichotomous items	81
▪ Introduction to the Rasch Model.....	81
▪ Item calibration.....	85
▪ Computation of a student's score.....	87
▪ Computation of a student's score for incomplete designs.....	91
▪ Optimal conditions for linking items.....	92
▪ Extension of the Rasch Model.....	93
Other item response theory models	94
Conclusion	94
 CHAPTER 6 PLAUSIBLE VALUES	 95
Individual estimates versus population estimates	96
The meaning of plausible values (PVs)	96
Comparison of the efficiency of WLEs, EAP estimates and PVs for the estimation of some population statistics	99
How to perform analyses with plausible values	102
Conclusion	103
 CHAPTER 7 COMPUTATION OF STANDARD ERRORS	 105
Introduction	106
The standard error on univariate statistics for numerical variables	106
The SAS® macro for computing the standard error on a mean	109
The standard error on percentages	112
The standard error on regression coefficients	115
The standard error on correlation coefficients	117
Conclusion	117
 CHAPTER 8 ANALYSES WITH PLAUSIBLE VALUES	 119
Introduction	120
Univariate statistics on plausible values	120
The standard error on percentages with PVs	123
The standard error on regression coefficients with PVs	123
The standard error on correlation coefficients with PVs	126
Correlation between two sets of plausible values	126
A fatal error shortcut	130
An unbiased shortcut	131
Conclusion	133
 CHAPTER 9 USE OF PROFICIENCY LEVELS	 135
Introduction	136
Generation of the proficiency levels	136
Other analyses with proficiency levels	141
Conclusion	143



CHAPTER 10 ANALYSES WITH SCHOOL-LEVEL VARIABLES	145
Introduction	146
Limits of the PISA school samples	147
Merging the school and student data files	148
Analyses of the school variables	148
Conclusion	150
CHAPTER 11 STANDARD ERROR ON A DIFFERENCE	151
Introduction	152
Statistical issues and computing standard errors on differences	152
The standard error on a difference without plausible values	154
The standard error on a difference with plausible values	159
Multiple comparisons	163
Conclusion	164
CHAPTER 12 OECD TOTAL AND OECD AVERAGE	167
Introduction	168
Recoding of the database to estimate the pooled OECD total and the pooled OECD average	170
Duplication of the data to avoid running the procedure three times	172
Comparisons between the pooled OECD total or pooled OECD average estimates and a country estimate	173
Comparisons between the arithmetic OECD total or arithmetic OECD average estimates and a country estimate	175
Conclusion	175
CHAPTER 13 TRENDS	177
Introduction	178
The computation of the standard error for trend indicators on variables other than performance	179
The computation of the standard error for trend indicators on performance variables	181
Conclusion	185
CHAPTER 14 STUDYING THE RELATIONSHIP BETWEEN STUDENT PERFORMANCE AND INDICES DERIVED FROM CONTEXTUAL QUESTIONNAIRES	187
Introduction	188
Analyses by quarters	188
The concept of relative risk	190
▪ Instability of the relative risk	191
▪ Computation of the relative risk	192
Effect size	195
Linear regression and residual analysis	197
▪ Independence of errors	197
Statistical procedure	200
Conclusion	201



CHAPTER 15 MULTILEVEL ANALYSES	203
Introduction	204
Two-level modelling with SAS®	206
▪ Decomposition of the variance in the empty model.....	206
▪ Models with only random intercepts.....	209
▪ Shrinkage factor.....	213
▪ Models with random intercepts and fixed slopes.....	213
▪ Models with random intercepts and random slopes.....	215
▪ Models with Level 2 independent variables.....	220
▪ Computation of final estimates and their respective standard errors.....	223
Three-level modelling	225
Limitations of the multilevel model in the PISA context	227
Conclusion	228
CHAPTER 16 PISA AND POLICY RELEVANCE – THREE EXAMPLES OF ANALYSES	231
Introduction	232
Example 1: Gender differences in performance	232
Example 2: Promoting socio-economic diversity within school?	236
Example 3: The influence of an educational system on the expected occupational status of students at age 30	242
Conclusion	246
CHAPTER 17 SAS® MACRO	247
Introduction	248
Structure of the SAS® Macro	248
REFERENCES	313
APPENDICES	315
Appendix 1 Three-level regression analysis.....	316
Appendix 2 PISA 2006 International database.....	324
Appendix 3 PISA 2006 Student questionnaire.....	333
Appendix 4 PISA 2006 Information communication technology (ICT) Questionnaire.....	342
Appendix 5 PISA 2006 School questionnaire.....	344
Appendix 6 PISA 2006 Parent questionnaire.....	351
Appendix 7 Codebook for PISA 2006 student questionnaire data file.....	355
Appendix 8 Codebook for PISA 2006 non-scored cognitive and embedded attitude items.....	399
Appendix 9 Codebook for PISA 2006 scored cognitive and embedded attitude items.....	419
Appendix 10 Codebook for PISA 2006 school questionnaire data file.....	431
Appendix 11 Codebook for PISA 2006 parents questionnaire data file.....	442
Appendix 12 PISA 2006 questionnaire indices.....	448



LIST OF BOXES

Box 2.1	WEIGHT statement in the proc means procedure	37
<hr/>		
Box 7.1	SAS® syntax for computing 81 means (e.g. PISA 2003).....	106
Box 7.2	SAS® syntax for computing the mean of HISEI and its standard error (e.g. PISA 2003).....	109
Box 7.3	SAS® syntax for computing the standard deviation of HISEI and its standard error by gender (e.g. PISA 2003).....	112
Box 7.4	SAS® syntax for computing the percentages and their standard errors for gender (e.g. PISA 2003).....	112
Box 7.5	SAS® syntax for computing the percentages and its standard errors for grades by gender (e.g. PISA 2003).....	114
Box 7.6	SAS® syntax for computing regression coefficients, R ² and its respective standard errors: Model 1 (e.g. PISA 2003).....	115
Box 7.7	SAS® syntax for computing regression coefficients, R ² and its respective standard errors: Model 2 (e.g. PISA 2003).....	116
Box 7.8	SAS® syntax for computing correlation coefficients and its standard errors (e.g. PISA 2003).....	117
<hr/>		
Box 8.1	SAS® syntax for computing the mean on the science scale by using the PROC_MEANS_NO_PV macro (e.g. PISA 2006).....	121
Box 8.2	SAS® syntax for computing the mean and its standard error on PVs (e.g. PISA 2006).....	122
Box 8.3	SAS® syntax for computing the standard deviation and its standard error on PVs by gender (e.g. PISA 2006).....	123
Box 8.4	SAS® syntax for computing regression coefficients and their standard errors on PVs by using the PROC_REG_NO_PV macro (e.g. PISA 2006).....	124
Box 8.5	SAS® syntax for running the simple linear regression macro with PVs (e.g. PISA 2006).....	125
Box 8.6	SAS® syntax for running the correlation macro with PVs (e.g. PISA 2006).....	126
Box 8.7	SAS® syntax for the computation of the correlation between mathematics/quantity and mathematics/space and shape by using the PROC_CORR_NO_PV macro (e.g. PISA 2003).....	129
<hr/>		
Box 9.1	SAS® syntax for generating the proficiency levels in science (e.g. PISA 2006).....	137
Box 9.2	SAS® syntax for computing the percentages of students by proficiency level in science and its standard errors by using the PROC_FREQ_NO_PV macro (e.g. PISA 2006).....	138
Box 9.3	SAS® syntax for computing the percentage of students by proficiency level in science and its standard errors by using the PROC_FREQ_PV macro (e.g. PISA 2006).....	140
Box 9.4	SAS® syntax for computing the percentage of students by proficiency level and its standard errors by gender (e.g. PISA 2006).....	140
Box 9.5	SAS® syntax for generating the proficiency levels in mathematics (e.g. PISA 2003).....	141
Box 9.6	SAS® syntax for computing the mean of self-efficacy in mathematics and its standard errors by proficiency level (e.g. PISA 2003).....	142
<hr/>		
Box 10.1	SAS® syntax for merging the student and school data files (e.g. PISA 2006).....	148
Box 10.2	Question on school location in PISA 2006.....	149
Box 10.3	SAS® syntax for computing the percentage of students and the average performance in science, by school location (e.g. PISA 2006).....	149
<hr/>		
Box 11.1	SAS® syntax for computing the mean of job expectations by gender (e.g. PISA 2003).....	154
Box 11.2	SAS® macro for computing standard errors on differences (e.g. PISA 2003).....	157

Box 11.3	Alternative SAS® macro for computing the standard error on a difference for a dichotomous variable (e.g. PISA 2003).....	158
Box 11.4	SAS® syntax for computing standard errors on differences which involve PVs (e.g. PISA 2003).....	160
Box 11.5	SAS® syntax for computing standard errors on differences that involve PVs (e.g. PISA 2006).....	162
<hr/>		
Box 12.1	SAS® syntax for computing the pooled OECD total for the mathematics performance by gender (e.g. PISA 2003).....	170
Box 12.2	SAS® syntax for the pooled OECD average for the mathematics performance by gender (e.g. PISA 2003).....	171
Box 12.3	SAS® syntax for the creation of a larger dataset that will allow the computation of the pooled OECD total and the pooled OECD average in one run (e.g. PISA 2003).....	172
<hr/>		
Box 14.1	SAS® syntax for the quarter analysis (e.g. PISA 2006).....	189
Box 14.2	SAS® syntax for computing the relative risk with five antecedent variables and five outcome variables (e.g. PISA 2006).....	193
Box 14.3	SAS® syntax for computing the relative risk with one antecedent variable and one outcome variable (e.g. PISA 2006).....	194
Box 14.4	SAS® syntax for computing the relative risk with one antecedent variable and five outcome variables (e.g. PISA 2006).....	194
Box 14.5	SAS® syntax for computing effect size (e.g. PISA 2006).....	196
Box 14.6	SAS® syntax for residual analyses (e.g. PISA 2003).....	200
<hr/>		
Box 15.1	Normalisation of the final student weights (e.g. PISA 2006).....	207
Box 15.2	SAS® syntax for the decomposition of the variance in student performance in science (e.g. PISA 2006).....	208
Box 15.3	SAS® syntax for normalising PISA 2006 final student weights with deletion of cases with missing values and syntax for variance decomposition (e.g. PISA 2006).....	211
Box 15.4	SAS® syntax for a multilevel regression model with random intercepts and fixed slopes (e.g. PISA 2006).....	214
Box 15.5	SAS® output for the multilevel model in Box 15.4.....	214
Box 15.6	SAS® syntax for a multilevel regression model (e.g. PISA 2006).....	216
Box 15.7	SAS® output for the multilevel model in Box 15.6.....	217
Box 15.8	SAS® output for the multilevel model with covariance between random parameters.....	218
Box 15.9	Interpretation of the within-school regression coefficient.....	220
Box 15.10	SAS® syntax for a multilevel regression model with a school-level variable (e.g. PISA 2006).....	221
Box 15.11	SAS® syntax for a multilevel regression model with interaction (e.g. PISA 2006).....	222
Box 15.12	SAS® output for the multilevel model in Box 15.11.....	222
Box 15.13	SAS® syntax for using the multilevel regression macro (e.g. PISA 2006).....	224
Box 15.14	SAS® syntax for normalising the weights for a three-level model (e.g. PISA 2006).....	226
<hr/>		
Box 16.1	SAS® syntax for testing the gender difference in standard deviations of reading performance (e.g. PISA 2000).....	233
Box 16.2	SAS® syntax for testing the gender difference in the 5th percentile of the reading performance (e.g. PISA 2006).....	235
Box 16.3	SAS® syntax for preparing a data file for the multilevel analysis.....	238



Box 16.4	SAS® syntax for running a preliminary multilevel analysis with one PV	239
Box 16.5	SAS® output for fixed parameters in the multilevel model.....	239
Box 16.6	SAS® syntax for running multilevel models with the PROC_MIXED_PV macro	242
<hr/>		
Box 17.1	SAS® macro of PROC_MEANS_NO_PV.sas.....	250
Box 17.2	SAS® macro of PROC_MEANS_PV.sas.....	253
Box 17.3	SAS® macro of PROC_FREQ_NO_PV.sas.....	256
Box 17.4	SAS® macro of PROC_FREQ_PV.sas.....	259
Box 17.5	SAS® macro of PROC_REG_NO_PV.sas.....	263
Box 17.6	SAS® macro of PROC_REG_PV.sas.....	266
Box 17.7	SAS® macro of PROC_CORR_NO_PV.sas.....	270
Box 17.8	SAS® macro of PROC_CORR_PV.sas.....	273
Box 17.9	SAS® macro of PROC_DIF_NO_PV.sas	276
Box 17.10	SAS® macro of PROC_DIF_PV.sas	279
Box 17.11	SAS® macro of QUARTILE_PV.sas	282
Box 17.12	SAS® macro of RELATIVE_RISK_NO_PV.sas.....	288
Box 17.13	SAS® macro of RELATIVE_RISK_PV.sas.....	291
Box 17.14	SAS® macro of EFFECT_SIZE_NO_PV.sas	296
Box 17.15	SAS® macro of EFFECT_SIZE_PV.sas.....	298
Box 17.16	SAS® macro of PROC_MIXED_NO_PV.sas.....	301
Box 17.17	SAS® macro of PROC_MIXED_PV.sas	306
<hr/>		
Box A1.1	Descriptive statistics of background and explanatory variables.....	318
Box A1.2	Background model for student performance.....	319
Box A1.3	Final net combined model for student performance.....	320
Box A1.4	Background model for the impact of socio-economic background.....	321
Box A1.5	Model of the impact of socio-economic background: “school resources” module	322
Box A1.6	Model of the impact of socio-economic background: “accountability practices” module	323
Box A1.7	Final combined model for the impact of socio-economic background.....	323

LIST OF FIGURES

Figure 1.1	Relationship between social and academic segregations.....	27
Figure 1.2	Relationship between social segregation and the correlation between science performance and student HISEI	27
Figure 1.3	Conceptual grid of variable types.....	29
Figure 1.4	Two-dimensional matrix with examples of variables collected or available from other sources	30
<hr/>		
Figure 2.1	Science mean performance in OECD countries (PISA 2006).....	38
Figure 2.2	Gender differences in reading in OECD countries (PISA 2000).....	38
Figure 2.3	Regression coefficient of ESCS on mathematic performance in OECD countries (PISA 2003).....	39
Figure 2.4	Design effect on the country mean estimates for science performance and for ESCS in OECD countries (PISA 2006)	42
Figure 2.5	Simple random sample and unbiased standard errors of ESCS on science performance in OECD countries (PISA 2006)	43



Figure 4.1	Distribution of the results of 36 students.....	60
Figure 4.2	Sampling variance distribution of the mean.....	62
Figure 5.1	Probability of success for two high jumpers by height (dichotomous).....	82
Figure 5.2	Probability of success for two high jumpers by height (continuous).....	83
Figure 5.3	Probability of success to an item of difficulty zero as a function of student ability.....	83
Figure 5.4	Student score and item difficulty distributions on a Rasch continuum.....	86
Figure 5.5	Response pattern probabilities for the response pattern (1, 1, 0, 0).....	88
Figure 5.6	Response pattern probabilities for a raw score of 1.....	89
Figure 5.7	Response pattern probabilities for a raw score of 2.....	90
Figure 5.8	Response pattern probabilities for a raw score of 3.....	90
Figure 5.9	Response pattern likelihood for an easy test and a difficult test.....	91
Figure 5.10	Rasch item anchoring.....	92
Figure 6.1	Living room length expressed in integers.....	96
Figure 6.2	Real length per reported length.....	97
Figure 6.3	A posterior distribution on a test of six items.....	98
Figure 6.4	EAP estimators.....	99
Figure 8.1	A two-dimensional distribution.....	127
Figure 8.2	Axes for two-dimensional normal distributions.....	127
Figure 13.1	Trend indicators in PISA 2000, PISA 2003 and PISA 2006.....	179
Figure 14.1	Percentage of schools by three school groups (PISA 2003).....	198
Figure 15.1	Simple linear regression analysis versus multilevel regression analysis.....	205
Figure 15.2	Graphical representation of the between-school variance reduction.....	215
Figure 15.3	A random multilevel model.....	216
Figure 15.4	Change in the between-school residual variance for a fixed and a random model.....	218
Figure 16.1	Relationship between the segregation index of students' expected occupational status and the segregation index of student performance in reading (PISA 2000).....	244
Figure 16.2	Relationship between the segregation index of students' expected occupational status and the correlation between HISEI and students' expected occupational status.....	245

LIST OF TABLES

Table 1.1	Participating countries/economies in PISA 2000, PISA 2003, PISA 2006 and PISA 2009.....	21
Table 1.2	Assessment domains covered by PISA 2000, PISA 2003 and PISA 2006.....	22
Table 1.3	Correlation between social inequities and segregations at schools for OECD countries.....	28
Table 1.4	Distribution of students per grade and per ISCED level in OECD countries (PISA 2006).....	31
Table 2.1	Design effect and type I errors.....	41
Table 2.2	Mean estimates and standard errors.....	45



Table 2.3	Standard deviation estimates and standard errors.....	45
Table 2.4	Correlation estimates and standard errors.....	45
Table 2.5	ESCS regression coefficient estimates and standard errors.....	46
<hr/>		
Table 3.1	Height and weight of ten persons	52
Table 3.2	Weighted and unweighted standard deviation estimate	52
Table 3.3	School, within-school, and final probability of selection and corresponding weights for a two-stage, simple random sample with the first-stage units being schools of equal size.....	54
Table 3.4	School, within-school, and final probability of selection and corresponding weights for a two-stage, simple random sample with the first-stage units being schools of unequal size	54
Table 3.5	School, within-school, and final probability of selection and corresponding weights for a simple and random sample of schools of unequal size (smaller schools)	55
Table 3.6	School, within-school, and final probability of selection and corresponding weights for a simple and random sample of schools of unequal size (larger schools)	55
Table 3.7	School, within-school, and final probability of selection and corresponding weights for PPS sample of schools of unequal size	56
Table 3.8	Selection of schools according to a PPS and systematic procedure.....	57
<hr/>		
Table 4.1	Description of the 630 possible samples of 2 students selected from 36 students, according to their mean.....	61
Table 4.2	Distribution of all possible samples with a mean between 8.32 and 11.68.....	63
Table 4.3	Distribution of the mean of all possible samples of 4 students out of a population of 36 students.....	64
Table 4.4	Between-school and within-school variances on the mathematics scale in PISA 2003.....	67
Table 4.5	Current status of sampling errors.....	67
Table 4.6	Between-school and within-school variances, number of participating schools and students in Denmark and Germany in PISA 2003	68
Table 4.7	The Jackknives replicates and sample means.....	70
Table 4.8	Values on variables X and Y for a sample of ten students.....	71
Table 4.9	Regression coefficients for each replicate sample.....	71
Table 4.10	The Jackknife replicates for unstratified two-stage sample designs.....	72
Table 4.11	The Jackknife replicates for stratified two-stage sample designs.....	73
Table 4.12	Replicates with the Balanced Repeated Replication method.....	74
Table 4.13	The Fay replicates	75
<hr/>		
Table 5.1	Probability of success when student ability equals item difficulty.....	84
Table 5.2	Probability of success when student ability is less than the item difficulty by 1 unit.....	84
Table 5.3	Probability of success when student ability is greater than the item difficulty by 1 unit	84
Table 5.4	Probability of success when student ability is less than the item difficulty by 2 units	85
Table 5.5	Probability of success when student ability is greater than the item difficulty by 2 units.....	85
Table 5.6	Possible response pattern for a test of four items.....	87
Table 5.7	Probability for the response pattern (1, 1, 0, 0) for three student abilities.....	87
Table 5.8	Probability for the response pattern (1, 0) for two students of different ability in an incomplete test design.....	91
Table 5.9	PISA 2003 test design	93

Table 6.1	Structure of the simulated data.....	100
Table 6.2	Means and variances for the latent variables and the different student ability estimators.....	100
Table 6.3	Percentiles for the latent variables and the different student ability estimators.....	101
Table 6.4	Correlation between HISEI, gender and the latent variable, the different student ability estimators.....	101
Table 6.5	Between- and within-school variances.....	102
<hr/>		
Table 7.1	HISEI mean estimates.....	107
Table 7.2	Squared differences between replicate estimates and the final estimate.....	108
Table 7.3	Output data file exercise1 from Box 7.2.....	111
Table 7.4	Available statistics with the PROC_MEANS_NO_PV macro.....	111
Table 7.5	Output data file exercise2 from Box 7.3.....	112
Table 7.6	Output data file exercise3 from Box 7.4.....	112
Table 7.7	Percentage of girls for the final and replicate weights and squared differences.....	113
Table 7.8	Output data file exercise4 from Box 7.5.....	114
Table 7.9	Output data file exercise5 from Box 7.6.....	115
Table 7.10	Output data file exercise6 from Box 7.7.....	116
Table 7.11	Output data file exercise6_criteria from Box 7.7.....	117
Table 7.12	Output data file exercise7 from Box 7.8.....	117
<hr/>		
Table 8.1	The 405 mean estimates.....	120
Table 8.2	Mean estimates and their respective sampling variances on the science scale for Belgium (PISA 2006).....	121
Table 8.3	Output data file exercise6 from Box 8.2.....	123
Table 8.4	Output data file exercise7 from Box 8.3.....	123
Table 8.5	The 450 regression coefficient estimates.....	125
Table 8.6	HISEI regression coefficient estimates and their respective sampling variance on the science scale in Belgium after accounting for gender (PISA 2006).....	125
Table 8.7	Output data file exercise8 from Box 8.5.....	125
Table 8.8	Output data file exercise9 from Box 8.6.....	126
Table 8.9	Correlation between the five plausible values for each domain, mathematics/quantity and mathematics/space and shape.....	128
Table 8.10	The five correlation estimates between mathematics/quantity and mathematics/space and shape and their respective sampling variance.....	129
Table 8.11	Standard deviations for mathematics scale using the correct method (plausible values) and by averaging the plausible values at the student level (pseudo-EAP) (PISA 2003).....	131
Table 8.12	Unbiased shortcut for a population estimate and its standard error.....	132
Table 8.13	Standard errors from the full and shortcut computation (PISA 2006).....	132
<hr/>		
Table 9.1	The 405 percentage estimates for a particular proficiency level.....	138
Table 9.2	Estimates and sampling variances per proficiency level in science for Germany (PISA 2006).....	139
Table 9.3	Final estimates of the percentage of students, per proficiency level, in science and its standard errors for Germany (PISA 2006).....	139
Table 9.4	Output data file exercise6 from Box 9.3.....	140
Table 9.5	Output data file exercise7 from Box 9.4.....	140
Table 9.6	Mean estimates and standard errors for self-efficacy in mathematics per proficiency level (PISA 2003).....	143
Table 9.7	Output data file exercise8 from Box 9.6.....	143



Table 10.1	Percentage of students per grade and ISCED level, by country (PISA 2006)	146
Table 10.2	Output data file exercise1 from Box 10.3	150
Table 10.3	Output data file exercise2 from Box 10.3	150
Table 11.1	Output data file exercise1 from Box 11.1	155
Table 11.2	Mean estimates for the final and 80 replicate weights by gender (PISA 2003)	155
Table 11.3	Difference in estimates for the final weight and 80 replicate weights between females and males (PISA 2003)	157
Table 11.4	Output data file exercise2 from Box 11.2	158
Table 11.5	Output data file exercise3 from Box 11.3	159
Table 11.6	Gender difference estimates and their respective sampling variances on the mathematics scale (PISA 2003)	159
Table 11.7	Output data file exercise4 from Box 11.4	160
Table 11.8	Gender differences on the mathematics scale, unbiased standard errors and biased standard errors (PISA 2003)	161
Table 11.9	Gender differences in mean science performance and in standard deviation for science performance (PISA 2006)	161
Table 11.10	Regression coefficient of HISEI on the science performance for different models (PISA 2006)	163
Table 11.11	Cross tabulation of the different probabilities	163
Table 12.1	Regression coefficients of the index of instrumental motivation in mathematics on mathematic performance in OECD countries (PISA 2003)	169
Table 12.2	Output data file exercise1 from Box 12.1	170
Table 12.3	Output data file exercise2 from Box 12.2	171
Table 12.4	Difference between the country mean scores in mathematics and the OECD total and average (PISA 2003)	174
Table 13.1	Trend indicators between PISA 2000 and PISA 2003 for HISEI, by country	180
Table 13.2	Linking error estimates	182
Table 13.3	Mean performance in reading by gender in Germany	184
Table 14.1	Distribution of the questionnaire index of cultural possession at home in Luxembourg (PISA 2006)	188
Table 14.2	Output data file exercise1 from Box 14.1	190
Table 14.3	Labels used in a two-way table	190
Table 14.4	Distribution of 100 students by parents' marital status and grade repetition	191
Table 14.5	Probabilities by parents' marital status and grade repetition	191
Table 14.6	Relative risk for different cutpoints	191
Table 14.7	Output data file exercise2 from Box 14.2	193
Table 14.8	Mean and standard deviation for the student performance in reading by gender, gender difference and effect size (PISA 2006)	195
Table 14.9	Output data file exercise4 from Box 14.5	197
Table 14.10	Output data file exercise5 from Box 14.5	197
Table 14.11	Mean of the residuals in mathematics performance for the bottom and top quarters of the PISA index of economic, social and cultural status, by school group (PISA 2003)	199

Table 15.1	Between- and within-school variance estimates and intraclass correlation (PISA 2006).....	209
Table 15.2	Output data file “ranparm1” from Box 15.3.....	212
Table 15.3	Output data file “fixparm3” from Box 15.6.....	217
Table 15.4	Output data file “ranparm3” from Box 15.6.....	217
Table 15.5	Variance/covariance estimates before and after centering.....	219
Table 15.6	Output data file of the fixed parameters file.....	221
Table 15.7	Average performance and percentage of students by student immigrant status and by type of school.....	223
Table 15.8	Variables for the four groups of students.....	223
Table 15.9	Comparison of the regression coefficient estimates and their standard errors in Belgium (PISA 2006).....	224
Table 15.10	Comparison of the variance estimates and their respective standard errors in Belgium (PISA 2006).....	225
Table 15.11	Three-level regression analyses.....	226
Table 16.1	Differences between males and females in the standard deviation of student performance (PISA 2000).....	234
Table 16.2	Distribution of the gender differences (males – females) in the standard deviation of the student performance.....	234
Table 16.3	Gender difference on the PISA combined reading scale for the 5 th , 10 th , 90 th and 95 th percentiles (PISA 2000).....	235
Table 16.4	Gender difference in the standard deviation for the two different item format scales in reading (PISA 2000).....	236
Table 16.5	Random and fixed parameters in the multilevel model with student and school socio-economic background.....	237
Table 16.6	Random and fixed parameters in the multilevel model with socio-economic background and grade retention at the student and school levels.....	241
Table 16.7	Segregation indices and correlation coefficients by country (PISA 2000).....	243
Table 16.8	Segregation indices and correlation coefficients by country (PISA 2006).....	244
Table 16.9	Country correlations (PISA 2000).....	245
Table 16.10	Country correlations (PISA 2006).....	246
Table 17.1	Synthesis of the 17 SAS® macros.....	249
Table A2.1	Cluster rotation design used to form test booklets for PISA 2006.....	324
Table A12.1	Mapping of ISCED to accumulated years of education.....	449
Table A12.2	ISCO major group white-collar/blue-collar classification.....	451
Table A12.3	ISCO occupation categories classified as science-related occupations.....	451
Table A12.4	Household possessions and home background indices.....	455
Table A12.5	Factor loadings and internal consistency of ESCS 2006 in OECD countries.....	465
Table A12.6	Factor loadings and internal consistency of ESCS 2006 in partner countries/economies.....	466



User's Guide

Preparation of data files

All data files (in text format) and the SAS® control files are available on the PISA website (www.pisa.oecd.org).

SAS® users

By running the SAS® control files, the PISA data files are created in the SAS® format. Before starting analysis, assigning the folder in which the data files are saved as a SAS® library.

For example, if the PISA 2000 data files are saved in the folder of "c:\pisa2000\data\", the PISA 2003 data files are in "c:\pisa2003\data\", and the PISA 2006 data files are in "c:\pisa2006\data\", the following commands need to be run to create SAS® libraries:

```
libname PISA2000 "c:\pisa2000\data\" ;  
libname PISA2003 "c:\pisa2003\data\" ;  
libname PISA2006 "c:\pisa2006\data\" ;  
run;
```

SAS® syntax and macros

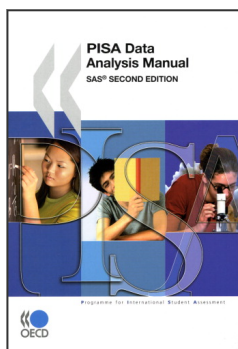
All syntaxes and macros in this manual can be copied from the PISA website (www.pisa.oecd.org). The 17 SAS® macros presented in Chapter 17 need to be saved under "c:\pisa\macro\", before starting analysis. Each chapter of the manual contains a complete set of syntaxes, which must be done sequentially, for all of them to run correctly, within the chapter.

Rounding of figures

In the tables and formulas, figures were rounded to a convenient number of decimal places, although calculations were always made with the full number of decimal places.

Country abbreviations used in this manual

AUS	Australia	FRA	France	MEX	Mexico
AUT	Austria	GBR	United Kingdom	NLD	Netherlands
BEL	Belgium	GRC	Greece	NOR	Norway
CAN	Canada	HUN	Hungary	NZL	New Zealand
CHE	Switzerland	IRL	Ireland	POL	Poland
CZE	Czech Republic	ISL	Iceland	PRT	Portugal
DEU	Germany	ITA	Italy	SVK	Slovak Republic
DNK	Denmark	JPN	Japan	SWE	Sweden
ESP	Spain	KOR	Korea	TUR	Turkey
FIN	Finland	LUX	Luxembourg	USA	United States



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